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# A Note on "The Optimal Stopping Time for Detecting Changes in Discrete Time Markov Processes" by Han and Tsung

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**Abstract:** We analyze the article by Han and Tsung (2009) "The Optimal Stopping Time for Detecting Changes in Discrete Time Markov Processes," and demonstrate that it is seriously flawed.

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#### 1. INTRODUCTION

Han and Tsung (2009) consider Markov observations  $X_1, X_2, ...$  with two possible stationary transition probabilities  $Pr(X_n = y | X_{n-1} = x) = p_i(x, y)$ , i = 0, 1. Up to some point in time t the samples follow  $p_0(x, y)$  and after t they switch to  $p_1(x, y)$ . The goal is to detect the change as soon as possible.

As a possible detection procedure, the authors propose the CUSUM stopping time  $T_p$ , which is defined with the help of the CUSUM statistic

$$V_n = \max\{V_{n-1}, 1\}l(X_{n-1} | X_n); \quad l(X_{n-1} | X_n) = \frac{p_1(X_{n-1}, X_n)}{p_0(X_{n-1}, X_n)}; \quad V_0 = 1,$$

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Address correspondence to G. V. Moustakides, Department of Electrical and Computer Engineering, University of Patras, Patras, Rio, 26500 Greece; E-mail: moustaki@upatras.gr or A. G. Tartakovsky, Department of Mathematics, University of Southern California, KAP-108, Los Angeles, CA 90089-2532, USA; E-mail: tartakov@math.usc.edu as  $T_p = \inf\{n \ge 1 : V_n \ge \mu\}$ . We emphasize that the threshold  $\mu$ , as in the case of the classical CUSUM, is considered *constant* and selected to satisfy the false alarm constraint  $E_0[T_p] = L$ , where L > 1 is some pre-specified constant.

As a performance measure, the authors use the Lorden (1971) criterion

$$D(T) = \sup_{m \ge 1} \operatorname{ess\,sup} \, E^{(m)}[(T - m + 1) \,|\, X_1, \dots, X_{m-1}]$$

In fact the definition of the Lorden measure given by the authors in Lemma 2.1 is incomplete, since it is missing the first supremum over  $m \ge 1$ . The goal is to minimize D(T) over all stopping times T that satisfy the false alarm constraint  $E_0[T] \ge L$ .

Following Moustakides (1986), the authors show in Lemma 2.1 that a lower bound can be found, namely

$$D(T) \ge \frac{E_0 \left[\sum_{n=0}^{T-1} \max\{V_n, 1\}\right]}{E_0 \left[\sum_{n=0}^{T-1} (1-V_n)^+\right]} = \overline{D}(T),$$
(1)

and then instead of optimizing D(T) over the class of stopping times  $E_0[T] \ge L$ , they optimize the lower bound  $\overline{D}(T)$  by minimizing the numerator and maximizing the denominator of  $\overline{D}(T)$  with the help of Theorem 2.1. In order for this double optimization to lead to the minimization of  $\overline{D}(T)$ , it is clear that *the same* (optimal) stopping time should be a solution to both optimization problems (for the numerator and the denominator).

## 2. DISCUSSION OF THE PROOF OF THEOREM 2.1

Theorem 2.1 in Han and Tsung (2009) claims optimality of CUSUM for the general constrained optimization problem

$$\sup_{T} E_0 \left[ \sum_{n=0}^{T-1} \varphi(V_n) \right] = E_0 \left[ \sum_{n=0}^{T_p-1} \varphi(V_n) \right]$$
(2)

for all T that satisfy  $E_0[T] = L$ , where  $\varphi(x)$  is a continuous nonincreasing function. This theorem is a key result for demonstrating the optimality of CUSUM in the sense of Lorden.

Consider the optimization problem in (2), which can be solved using the Optimal Stopping Theory in Shiryaev (1978). Note that in order to apply Shiryaev's results the decision statistic has to be a *Markov process*, and for the CUSUM procedure (with a constant threshold) to be optimal the CUSUM statistic  $\{V_n\}_{n\geq 1}$  should be a homogeneous Markov process, as in the i.i.d. case. However, for the Markov model considered the CUSUM statistic is NOT Markov. In order to obtain a Markov process we need to combine  $V_n$  with the observation  $X_n$ . In other words, the two-dimensional process  $\{(V_n, X_n)\}_{n\geq 1}$  has the desired Markov property.

To optimize (2) we reduce it into an unconstrained problem and consider

$$U(v, x) = \sup_{T} E_0 \left[ \sum_{n=0}^{T-1} \{ \varphi(V_n) - \lambda \} \right]$$

where  $V_0 = v$ ,  $X_0 = x$ . The optimum cost function U(v, x) depends, of course, on the initial state of the Markov process  $\{(V_n, X_n)\}$ . By Theorem 2.23 of Shiryaev (1978), U(v, x) satisfies the following equation

$$U(v, x) = \max\{0, \varphi(v) - \lambda + E_0[U(V_1, X_1) | V_0 = v, X_0 = x]\}.$$

It is clear that  $E_0[U(V_1, X_1) | V_0 = v, X_0 = x] = \overline{U}(v, x)$  is a function of both quantities v, x. Consequently, finding the optimum stopping boundary requires solving the equation  $\varphi(v) - \lambda + \overline{U}(v, x) = 0$ , which has a solution of the form  $v = \mu(x)$ , but by no means a constant boundary  $v = \mu$ . In other words, the optimal threshold is a function of x, and the CUSUM test needs to be modified as  $T_p = \inf\{n : V_n \ge \mu(X_n)\}$ . Let us call this CUSUM test the (V, X)-CUSUM, to emphasize the difference with the traditional CUSUM.

One might suggest that we can simply adopt this modification for the CUSUM procedure, i.e., introduce a variable threshold, and this will be sufficient to guarantee the desired optimality. However, we note that the threshold  $\mu(x)$  depends in general on the form of the function  $\varphi(x)$ . Since Theorem 2.1 is used to optimize, separately, the numerator and the denominator in  $\overline{D}(T)$  defined in (1), there is absolutely no guarantee that the optimum threshold will turn out to be the same in both optimization problems.

Let us disregard even this latter complication and assume that there exists a unique variable threshold  $\mu(x)$  for CUSUM that optimizes the lower bound  $\overline{D}(T)$ . Does this necessarily imply optimization of the original Lorden measure D(T) as in the case of i.i.d. observations?

#### 3. DISCUSSION OF THE OPTIMIZATION OF LORDEN'S MEASURE

Note from (1) that the Lorden measure D(T) is lower-bounded by  $\overline{D}(T)$ . By our previous assumption, there exists a variable threshold CUSUM version  $T_p$ that optimizes  $\overline{D}(T)$ . In order for this to also imply optimization of the original Lorden measure D(T), the equality  $D(T_p) = \overline{D}(T_p)$  should hold. In other words, for (V, X)-CUSUM, Lorden's measure and the lower bound must coincide. Indeed, if this were the case, then from (1), by minimizing over T both sides, we could argue that  $\inf_T D(T) \ge \inf_T \overline{D}(T) = \overline{D}(T_p) = D(T_p)$ . This would also imply  $\inf_T D(T) \ge$  $D(T_p)$ , i.e., optimality of the (V, X)-CUSUM procedure in Lorden's sense. It is therefore imperative to show that for (V, X)-CUSUM  $D(T_p)$  and  $\overline{D}(T_p)$  are equal. Unfortunately this is not the case, as we briefly explain next.

Note that due to the Markov structure of the sequence  $\{(V_n, X_n)\}_{n>1}$ 

$$E^{(m)}[(T_p - m + 1)^+ | X_0, \dots, X_{m-1}] = \psi(V_{m-1}, X_{m-1})$$

for some suitable function  $\psi(v, x)$ ; where, we recall,  $E^{(m)}[\cdot]$  denotes expectation with respect to the probability measure induced by the change occurring at time *m*. Since for a given  $X_{m-1}$  the statistic  $V_n$ ,  $n \ge m$ , is path-wise increasing in  $V_{m-1}$ , we conclude that the detection delay  $\psi(V_{m-1}, X_{m-1})$  is decreasing in  $V_{m-1}$ , implying that  $\psi(V_{m-1}, X_{m-1}) \le \psi(1, X_{m-1})$ . In other words, for a given  $X_{m-1}$  the conditional expectation is maximized when we restart the CUSUM process. Unlike the i.i.d. case  $\psi(1, X_{m-1})$  does not correspond to the worst performance of CUSUM, since we need to further maximize over  $X_{m-1}$ . Indeed, note that

ess sup 
$$E^{(m)}[(T_p - m + 1)^+ | X_0, \dots, X_{m-1}] = \operatorname{ess sup} \psi(V_{m-1}, X_{m-1}) = \sup_{x} \psi(1, x).$$

This has a grave consequence on the relationship between  $D(T_p)$  and  $\overline{D}(T_p)$ . The lower bound  $\overline{D}(T_p)$  is obtained by summing the following quantities (see Lemma 3 in Moustakides, 1986)

$$\begin{split} E_0 \Big[ E^{(m)} \big[ (T_p - m + 1)^+ \, | \, X_0, \dots, X_{m-1} \big] (1 - V_{m-1})^+ \Big] &= E_0 \big[ \psi(V_{m-1}, X_{m-1}) (1 - V_{m-1})^+ \big] \\ &= E_0 \big[ \psi(1, X_{m-1}) (1 - V_{m-1})^+ \big] \\ &\leq \Big\{ \sup_x \psi(1, x) \Big\} E_0 \big[ (1 - V_{m-1})^+ \big]. \end{split}$$

The second equality is true because every time  $V_n \leq 1$  the CUSUM process restarts from 1. Note now that the last inequality *is strict* whenever  $\psi(1, x)$  is not a constant. While equality holds for i.i.d. observations (since  $\psi(V, x) = \psi(V)$ ), the same property is NOT valid for Markov observations. Thus the previous inequality is strict yielding the strict inequality  $D(T_p) > \overline{D}(T_p)$  between Lorden's measure and its lower bound. Consequently, even if the (V, X)-CUSUM test optimizes  $\overline{D}(T)$ , this does not necessarily imply that it optimizes the original Lorden measure D(T).

# 4. CONCLUSION

Although we are sure that the conventional CUSUM is not optimal in the Markov case, we would like to stress that in this note we did not rigorously prove that Han and Tsung's claim regarding optimality of CUSUM is wrong. Rather we showed that crucial parts of their proof are problematic. We believe that finding the optimal detection procedure in the Markov case requires fresh ideas that go far beyond immediate extensions of existing methodologies. Furthermore, if a CUSUM-like procedure turns out to be optimal, we believe that not only the stopping threshold, but also the restarting barrier (which is 1 in the classical case) will be functions of the current observation  $X_n$ .

## REFERENCES

- Han, D. and Tsung, F. (2009). The Optimal Stopping Time for Detecting Changes in Discrete Time Markov Processes, Sequential Analysis 28: 115–135.
- Lorden, G. (1971). Procedures for Reacting to a Change in Distribution, Annals of Mathematical Statistics 42: 1897–1908.
- Moustakides, G. V. (1986). Optimal Stopping Times for Detecting Changes in Distributions, Annals of Statistics 14: 1379–1387.

Shiryaev, A. N. (1978). Optimal Stopping Rules, New York: Springer-Verlag.