Optimum Sequential Procedures

for

Detecting Changes in Processes

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Outline

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- Overview of existing results
- Lorden's criterion and the CUSUM test
- A modified Lorden criterion
- Optimality of CUSUM for Itô processes

The Change Detection (Disorder) Problem

We are observing sequentially a process ξ_t with the following statistics

$$\begin{array}{ll} \xi_t & \sim & \mathbb{P}_{\infty} & \quad \text{for } 0 \leq t \leq \tau \\ & \sim & \mathbb{P}_0 & \quad \text{for } \tau < t \end{array}$$

- Change time τ : deterministic (but unknown) or random.
- Probability measures $\mathbb{P}_{\infty}, \mathbb{P}_{0}$: known.

Detect the change "as soon as possible".

Applications include: systems monitoring; quality control; financial decision making; remote sensing (radar, sonar, seismology); occurrence of industrial accidents; speech/image/video segmentation; etc.

The observation process ξ_t is available sequentially; this can be expressed through the filtration:

 $\mathcal{F}_t = \sigma\{\xi_s : 0 \le s \le t\}.$

For detecting the change we are interested in sequential schemes.

Any sequential detection scheme can be represented by a stopping time T (the time we stop and declare that the change took place).

The stopping time T is adapted to \mathcal{F}_t .

In other words, at every time instant t we perform a test (whether to stop and declare a change or continue sampling) using only the available information up to time t.

Overview of Existing Results

- \mathbb{P}_{τ} : the probability measure induced, when the change takes place at time τ .
- $\mathbb{E}_{\tau}[\cdot]$: the corresponding expectation.
- \mathbb{P}_{∞} : all data under nominal régime.
- \mathbb{P}_0 : all data under alternative régime.

Optimality Criteria

- They are basically comprised of two parts:
- The first measures the detection delay
- The second the frequency of false alarms

Possible approaches are Baysian and Min-max.

Bayesian Approach (Shiryayev):

 τ is random and exponentially distributed

$$\inf_{T} \{ c \mathbb{E}[(T-\tau)^+] + \mathbb{P}[T < \tau] \}$$

The Shiryayev test consists in computing the statistics $\pi_t = \mathbb{P}[\tau \le t | \mathcal{F}_t]$; and stop when

$$T_S = \inf_t \{t : \pi_t \ge \nu\}.$$

 T_S is optimum (Shiryayev 1978):

- In discrete time: when ξ_n is i.i.d. before and after the change.
- In continuous time: when ξ_t is a Brownian Motion with constant drift before and after the change.

Min-Max Approach (Shiryayev-Roberts-Pollak): τ is deterministic and unknown

$$\inf_T \sup_{\tau} \mathbb{E}_{\tau}[(T-\tau)^+ | T > \tau]; \text{ subject } \mathbb{E}_{\infty}[T] \ge \gamma.$$

Optimality results exists only for discrete time when ξ_n is i.i.d. before and after the change. Specifically if we define the statistics

$$S_n = (S_{n-1} + 1) \frac{f_0(\xi_n)}{f_\infty(\xi_n)},$$

where $f_{\infty}(\cdot), f_0(\cdot)$ the common pdf of the data before and after the change then (Yakir 1997) the stopping time

$$T_{SRP} = \inf_n \{ n : S_n \ge \nu \}$$

is optimum.

Lorden's Criterion and the CUSUM Test

An alternative min-max approach consists in defining the following performance measure (Lorden 1971)

$$J(T) = \sup_{\tau} \operatorname{essup} \mathbb{E}_{\tau}[(T-\tau)^{+} | \mathcal{F}_{\tau}]$$

and solve the min-max problem

$$\inf_{T} J(T); \text{ subject to } \mathbb{E}_{\infty}[T] \geq \gamma.$$

The test closely related to Lorden's criterion and being to most popular one used in practice is the **Cumulative Sum** (CUSUM) test.

Define the CUSUM statistics y_t as follows:

$$u_t = \log\left(\frac{d\mathbb{P}_0}{d\mathbb{P}_\infty}(\mathcal{F}_t)\right); \quad m_t = \inf_{0 \le s \le t} u_s$$
$$y_t = u_t - m_t.$$

The CUSUM stopping time (Page 1954):

$$T_C = \inf_t \{t : y_t \ge \nu\}.$$

Optimality results:

- Discrete time: when ξ_n is i.i.d. before and after the change (Moustakides 1986, Ritov 1990).
- Continuous time: when ξ_t is a Brownian Motion with constant drift before and after the change (Shiryayev 1996, Beibel 1996).

A modified Lorden criterion

Our goal is to extend the optimality of CUSUM to Itô processes. For this it will be necessary to modify Lorden's criterion using the Kullback-Leibler Divergence (KLD).

Similar extension was proposed for the SPRT by Liptser and Shiryayev (1978).

Consider the process ξ_t

$$d\xi_t = \begin{cases} dw_t, & 0 \le t \le \tau \\ \alpha_t dt + dw_t, & \tau < t \end{cases}$$

where w_t is a standard Brownian motion with respect to $\mathcal{F}_t = \sigma(\xi_s; 0 \le s \le t); \alpha_t$ is adapted to \mathcal{F}_t and τ denotes the time of change.

To ξ_t we correspond the process u_t defined by

$$du_t = \alpha_t d\xi_t - 0.5\alpha_t^2 dt$$

which we like to play the role of the log-likelihood ratio $u_t = \log(d\mathbb{P}_0/d\mathbb{P}_\infty(\mathcal{F}_t))$. We therefore need to impose the following conditions:

1.
$$\mathbb{P}_0\left[\int_0^t \alpha_s^2 ds < \infty\right] = \mathbb{P}_\infty\left[\int_0^t \alpha_s^2 ds < \infty\right] = 1$$

2. A "Novikov" condition, i.e. $\mathbb{E}_{\infty}[\exp(\int_{t_{n-1}}^{t_n} \alpha_s^2 ds)] < \infty$ where t_n strictly increasing with $t_n \to \infty$.

3.
$$\mathbb{P}_0\left[\int_0^\infty \alpha_s^2 ds = \infty\right] = \mathbb{P}_\infty\left[\int_0^\infty \alpha_s^2 ds = \infty\right] = 1$$

From conditions 1 & 2 we have validity of Girsanov's theorem, therefore

$$\frac{d\mathbb{P}_0}{d\mathbb{P}_{\infty}}(\mathcal{F}_t) = e^{u_t}; \quad \frac{d\mathbb{P}_{\tau}}{d\mathbb{P}_{\infty}}(\mathcal{F}_t) = e^{u_t - u_{\tau}}.$$

Furthermore for the KLD we can write

$$\begin{split} \mathbb{E}_{\tau} \left[\log \left(\frac{d\mathbb{P}_{\tau}}{d\mathbb{P}_{\infty}} (\mathcal{F}_{t}) \right) \middle| \mathcal{F}_{\tau} \right] \\ &= \mathbb{E}_{\tau} \left[\int_{\tau}^{t} \alpha_{s} dw_{s} + \int_{\tau}^{t} \frac{1}{2} \alpha_{s}^{2} ds \middle| \mathcal{F}_{\tau} \right] \\ &= \mathbb{E}_{\tau} \left[\int_{\tau}^{t} \frac{1}{2} \alpha_{s}^{2} ds \middle| \mathcal{F}_{\tau} \right], \text{ for } 0 \leq \tau \leq t < \infty, \end{split}$$

This suggests the following modification in Lorden's criterion

$$J(T) = \sup_{\tau \in [0,\infty)} \operatorname{essup} \mathbb{E}_{\tau} \left[\mathbb{1}_{\{T > \tau\}} \int_{\tau}^{T} \frac{1}{2} \alpha_{t}^{2} dt \, \Big| \, \mathcal{F}_{\tau} \right],$$

and the corresponding min-max optimization

$$\inf_{T} J(T); \text{ subject } \mathbb{E}_{\infty} \left[\int_{0}^{T} \frac{1}{2} \alpha_{t}^{2} dt \right] \geq \gamma.$$

The original and the modified criterion coincide when ξ_t is a Brownian motion with constant drift.

Let us form the CUSUM statistics y_t for the Itô process

$$du_t = \alpha_t d\xi_t - 0.5\alpha_t^2 dt$$
$$m_t = \inf_{\substack{0 \le s \le t}} u_s$$
$$y_t = u_t - m_t$$

and the optimum CUSUM test is

$$T_C = \inf_t \{t : y_t \ge \nu\}; \text{ where } \mathbb{E}_{\infty} \left[\int_0^{T_C} \frac{1}{2} \alpha_t^2 dt \right] = \gamma.$$

Since y_t has continuous paths we conclude that when the CUSUM test stops we will have: $y_{T_C} = \nu$.

Optimality of CUSUM for Itô processes



 $u_t \ge m_t$ therefore $y_t = u_t - m_t \ge 0$. m_t is nonincreasing and $dm_t \ne 0$ only when $u_t = m_t$ or $y_t = 0$. If f(y) continuous; f(0) = 0, then $\int_0^\infty f(y_t) dm_t = 0$. If f(y) is a twice continuously differentiable function with f'(0) = 0, using standard Itô calculus we have

$$df(y_t) = f'(y_t)(du_t - dm_t) + 0.5\alpha_t^2 f''(y_t)dt = f'(y_t)du_t + 0.5\alpha_t^2 f''(y_t)dt$$

Theorem 1:
$$T_C$$
 is a.s. finite and

$$\mathbb{E}_{\tau} \left[\mathbbm{1}_{\{T_C > \tau\}} \int_{\tau}^{T_C} \frac{1}{2} \alpha_t^2 dt \, \middle| \, \mathcal{F}_{\tau} \right] = [g(\nu) - g(y_{\tau})] \mathbbm{1}_{\{T_C > \tau\}}$$

$$\mathbb{E}_{\infty} \left[\mathbbm{1}_{\{T_C > \tau\}} \int_{\tau}^{T_C} \frac{1}{2} \alpha_t^2 dt \, \middle| \, \mathcal{F}_{\tau} \right] = [h(\nu) - h(y_{\tau})] \mathbbm{1}_{\{T_C > \tau\}}.$$
where

$$g(y) = y + e^{-y} - 1;$$
 $h(y) = e^{y} - y - 1.$

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Since g(y), h(y) are increasing and strictly convex with g(0) = h(0) = 0, we now conclude

$$J(T_C) = \sup_{\tau} \operatorname{essup} \mathbb{E}_{\tau} \left[0.5 \int_{\tau}^{T_C} \alpha_s^2 ds | \mathcal{F}_{\tau} \right]$$
$$= \sup_{\tau} \operatorname{essup} [g(\nu) - g(y_{\tau})] \mathbb{1}_{\{T_C > \tau\}}$$
$$= g(\nu) - g(0) = g(\nu)$$

Similarly

$$\mathbb{E}_{\infty}\left[\int_{0}^{T_{C}} \alpha_{s}^{2} ds\right] = h(\nu) - h(0) = h(\nu) = \gamma.$$

The threshold can thus be computed: $e^{\nu} - \nu - 1 = \gamma$.

Using again standard Itô calculus we have the following generalization of Theorem 1.

Corollary:

$$\mathbb{E}_{\tau} \left[\int_{\tau}^{T} \frac{1}{2} \alpha_{t}^{2} dt \, \middle| \, \mathcal{F}_{\tau} \right] = \mathbb{E}_{\tau} \left[g(y_{T}) - g(y_{\tau}) \middle| \mathcal{F}_{\tau} \right] \mathbb{1}_{\{T > \tau\}}$$
$$\mathbb{E}_{\infty} \left[\int_{\tau}^{T} \frac{1}{2} \alpha_{t}^{2} dt \, \middle| \, \mathcal{F}_{\tau} \right] = \mathbb{E}_{\infty} \left[h(y_{T}) - h(y_{\tau}) \middle| \mathcal{F}_{\tau} \right] \mathbb{1}_{\{T > \tau\}}$$

where T stopping time.

Remark 1: The false alarm constraint can be written as

$$\mathbb{E}_{\infty}\left[\int_{0}^{T} \frac{1}{2}\alpha_{t}^{2}dt\right] = \mathbb{E}_{\infty}[h(y_{T})] \ge \gamma$$

Remark 2: We can limit ourselves to stopping times that satisfy the false alarm constraint with equality, that is,

$$\mathbb{E}_{\infty}\left[\int_{0}^{T} \frac{1}{2}\alpha_{t}^{2}dt\right] = \mathbb{E}_{\infty}[h(y_{T})] = \gamma = h(\nu).$$

Remark 3: The modified performance measure J(T) can be suitably lower bounded as follows

$$J(T) = \sup_{\tau} \operatorname{essup} \mathbb{E}_{\tau} \left[\mathbb{1}_{\{T > \tau\}} \int_{\tau}^{T} \frac{1}{2} \alpha_{t}^{2} dt \, \Big| \, \mathcal{F}_{\tau} \right]$$
$$\geq \frac{\mathbb{E}_{\infty} \left[e^{y_{T}} g(y_{T}) \right]}{\mathbb{E}_{\infty} \left[e^{y_{T}} \right]}.$$

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Theorem 2: Any stopping time T that satisfies the false alarm constraint with equality has a performance measure J(T) that is no less than $J(T_C) = g(\nu)$.

Proof: To show $J(T) \geq g(\nu),$ since

$$J(T) \ge \frac{\mathbb{E}_{\infty}\left[e^{y_T}g(y_T)\right]}{\mathbb{E}_{\infty}\left[e^{y_T}\right]},$$

it is sufficient to show that

$$\frac{\mathbb{E}_{\infty}\left[e^{y_T}g(y_T)\right]}{\mathbb{E}_{\infty}\left[e^{y_T}\right]} \ge g(\nu)$$

or equivalently:
$$\mathbb{E}_{\infty}\left[e^{y_T}\left\{g(y_T) - g(\nu)\right\}\right] \geq 0$$

We recall that we consider stopping times with

$$\mathbb{E}_{\infty}\left[\int_{0}^{T} \frac{1}{2}\alpha_{t}^{2}dt\right] = \mathbb{E}_{\infty}[h(y_{T})] = \gamma = h(\nu),$$

therefore the inequality we like to prove is equivalent to

$$\mathbb{E}_{\infty} \left[e^{y_T} \{ g(y_T) - g(\nu) \} + h(\nu) - h(y_T) \right] \ge 0.$$

The function

$$p(y) = e^{y} \{ g(y) - g(\nu) \} + h(\nu) - h(y)$$

for $y \geq 0$, can be shown to exhibit a global minimum at $y = \nu$



Because $p(\nu) = 0$, we conclude that $p(y) \ge 0$, thus $\mathbb{E}_{\infty}[p(y_T)] \ge 0$

with equality iff $y_T = \nu$ (i.e. the CUSUM stopping time).

Conclusion

- We introduced a modification of Lorden's criterion based on the Kullback-Leibler Divergence for the problem of detecting changes in Itô processes.
- With the help of the new criterion we introduced a constrained min-max optimization problem that defines the optimum sequential scheme for the change detection problem.
- We demonstrated that the CUSUM test is the solution to the above optimization problem.