

From Statistical Estimation and Generative Modeling to Inverse Problems

LECTURE 3: STATISTICAL ESTIMATION

Prologue

In our three lectures we focus on the following subjects:

- Statistical Estimation methods (Parameter Estimation)
Measure X but would like to know (estimate) Z
- Random data are described by probability densities
Not convenient for modern datasets: Usually multi-dimensional but reside on low dimensional surfaces (manifolds)
Alternative description with Generative Models
- Modification of Statistical Estimation methods to accept generative models in place of probability densities
Application to inverse problems (mostly from Computer Vision)

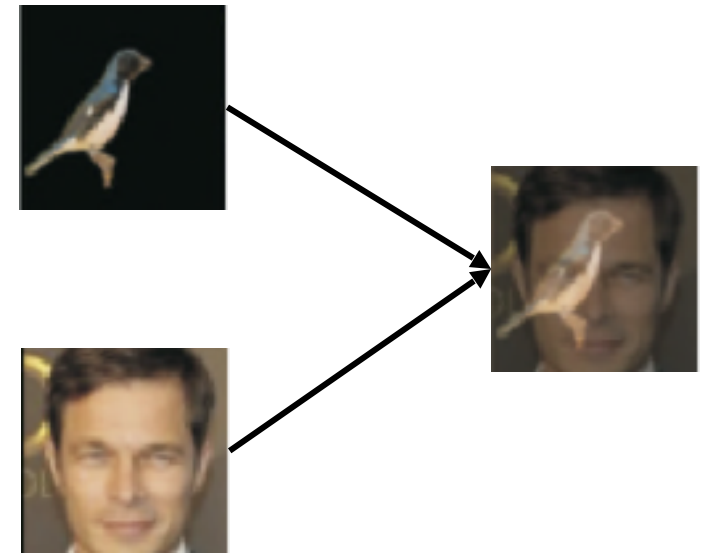
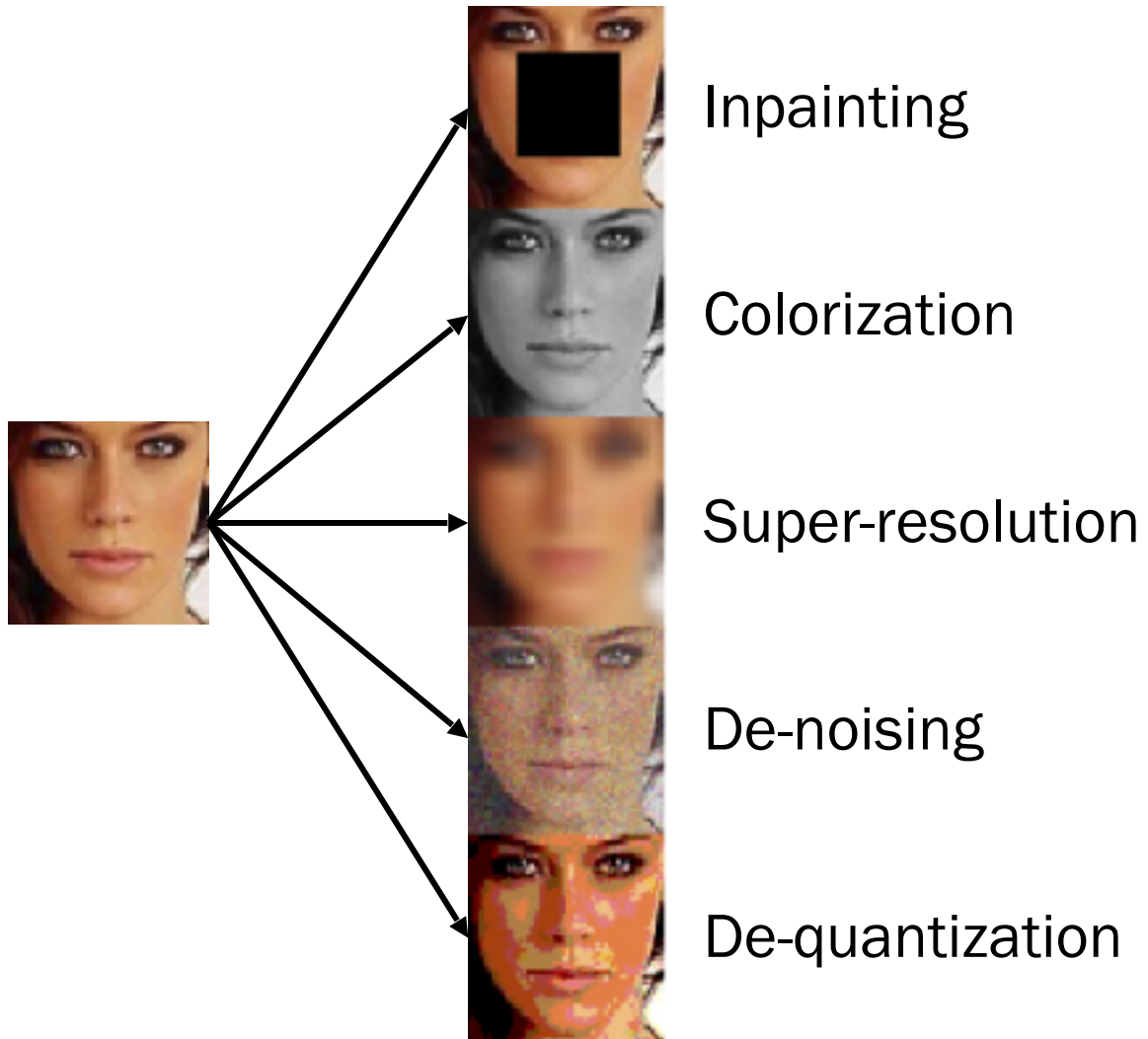


Image Separation

Statistical Estimation

- Problem Definition
- Bayesian Estimation
 - Optimum solution for general cost function
 - Solution for the MSE, MAE, AP
- Non-Bayesian Estimation
 - The ML estimator
 - Asymptotic optimality

Problem Definition

We are given vector X (measurements) but interested in vector Z !

Can we estimate Z from X ?

Z and X must be **related** otherwise impossible

The relationship can be “soft” (not deterministic) and expressed with a joint probability density $f(X, Z)$

So Z and X are **random** and have a **random** relationship captured by $f(X, Z)$

The density $f(X, Z)$ also expresses prior knowledge about each individual random quantity Z and X

Prior knowledge about individual X and Z is captured by the marginal densities $g(X)$ and $h(Z)$ where

$$g(X) = \int f(X, Z) dZ, \quad h(Z) = \int f(X, Z) dX$$

When measurements X become available, knowledge about Z changes to

$$f(Z|X) = \frac{f(X, Z)}{g(X)} = \frac{f(X, Z)}{\int f(X, Z) dZ}$$

If X and Z independent then $f(X, Z) = g(X)h(Z)$ and

$$f(Z|X) = \frac{f(X, Z)}{g(X)} = \frac{g(X)h(Z)}{g(X)} = h(Z)$$

X has no information about Z

What is an estimator ?

Given $f(X, Z)$ and measurements X suppose we select “estimate” \hat{Z}

If measurements X do not change no reason to change \hat{Z}

Estimate $\hat{Z} = \hat{Z}(X)$ is **any deterministic function** of X

There are “good” estimators and there are “bad” estimators

Statistical Estimation theory identifies the “best”

Bayesian Estimation

There is cost function $C(\cdot, \cdot)$ such that if true value is Z and estimate \hat{Z} then cost is $C(\hat{Z}, Z)$

Interested in **average cost**

$$\mathcal{C}(\hat{Z}) = \mathbb{E}_{X,Z} [C(\hat{Z}(X), Z)] = \iint C(\hat{Z}(X), Z) f(X, Z) dX dZ$$

Depends on the estimator function $\hat{Z}(X)$

We like to **minimize** it

$$\min_{\hat{Z}} \mathcal{C}(\hat{Z}) = \min_{\hat{Z}} \iint C(\hat{Z}(X), Z) f(X, Z) dX dZ$$

Prior knowledge about X and Z

$$g(X) = \int f(X, Z) dZ, \quad h(Z) = \int f(X, Z) dX$$

After we take measurements X our knowledge about Z becomes

$$f(Z|X) = \frac{f(X, Z)}{g(X)} = \frac{f(X, Z)}{\int f(X, Z) dZ}$$

Posterior probability
density

and we can write $f(X, Z) = f(Z|X)g(X)$

$$\mathcal{C}(\hat{Z}) = \int \left(\int \mathcal{C}(\hat{Z}(X), Z) f(Z|X) dZ \right) g(X) dX$$

call $G(\hat{Z}(X), X)$

$$G(U, X) = \int C(U, Z) f(Z|X) dZ$$

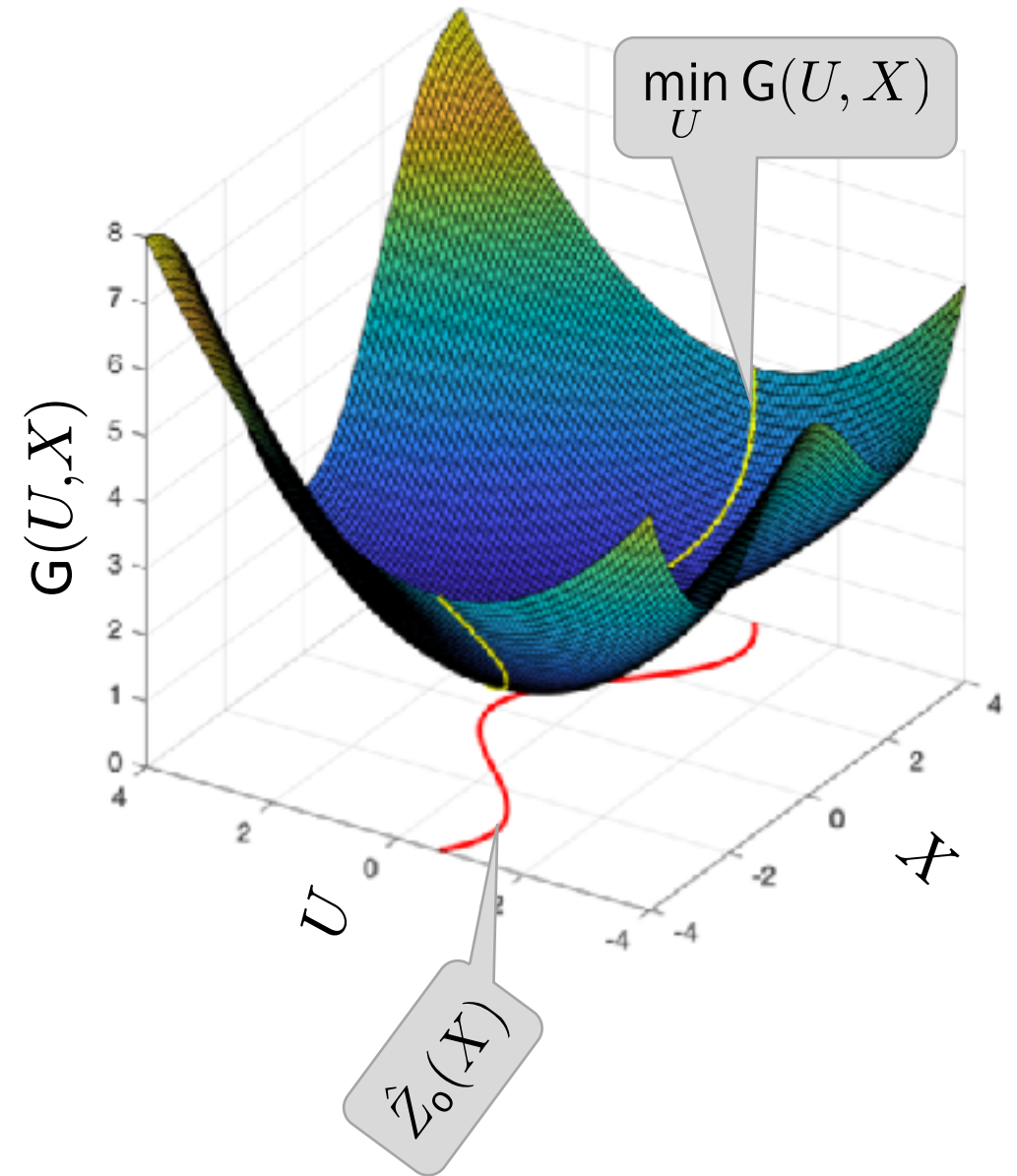
$$\mathcal{L}(\hat{Z}) = \int G(\hat{Z}(X), X) g(X) dX$$

For each X perform minimization

$$\min_U G(U, X) = \Phi(X)$$

$$\hat{Z}_0(X) = \arg \min_U G(U, X)$$

$$G(\hat{Z}_0(X), X) = \min_U G(U, X) = \Phi(X)$$



$$\min_U \mathbf{G}(U, X) = \Phi(X) \Rightarrow \mathbf{G}(??, X) \geq \Phi(X) \Rightarrow \mathbf{G}(\hat{Z}(X), X) \geq \Phi(X)$$

$$\mathcal{L}(\hat{Z}) = \int \mathbf{G}(\hat{Z}(X), X) g(X) dX \geq \int \Phi(X) g(X) dX$$

Lower bound

If estimator attains lower bound then **optimum**

Consider $\hat{Z}_o(X) = \arg \min_U \mathbf{G}(U, X)$

$$\mathcal{L}(\hat{Z}_o) = \int \underbrace{\mathbf{G}(\hat{Z}_o(X), X)}_{= \Phi(X)} g(X) dX = \int \Phi(X) g(X) dX$$

$\hat{Z}_o(X) = \arg \min_U \mathbf{G}(U, X)$ is optimum

Examples

Minimum Mean Square Error (MMSE): $C(\hat{Z}, Z) = \|\hat{Z} - Z\|^2$

$$G(U, X) = \int C(U, Z) f(Z|X) dZ = \int \|U - Z\|^2 f(Z|X) dZ$$

Must compute $\min_U G(U, X)$

$$\nabla_U G(U, X) = 0 \Rightarrow \int \nabla_U \|U - Z\|^2 f(Z|X) dZ = 0$$

$$\int 2(U - Z) f(Z|X) dZ = 0 \Rightarrow U = \int Z f(Z|X) dZ$$

$$\hat{Z}_{\text{MMSE}}(X) = \mathbb{E}[Z|X] = \int Z f(Z|X) dZ = \frac{\int Z f(X, Z) dZ}{\int f(X, Z) dZ}$$

Minimum Mean Absolute Error (MMAE): $C(\hat{Z}, Z) = |\hat{z}_1 - z_1| + \cdots + |\hat{z}_L - z_L|$

$$\hat{Z} = [\hat{z}_1 \cdots \hat{z}_L]^\top, \quad Z = [z_1 \cdots z_L]^\top$$

Treat each coordinate separately

Minimize with respect to u : $G(u, X) = \int |u - z| f(z|X) dz$

$$\frac{\partial G(u, X)}{\partial u} = 0 \Rightarrow \int \frac{\partial |u - z|}{\partial u} f(z|X) dz = \int \text{sign}(u - z) f(z|X) dz = 0$$

$$\Rightarrow \int_{-\infty}^u f(z|X) dz = \int_u^{\infty} f(z|X) dz \Rightarrow 2 \int_{-\infty}^u f(z|X) dz = 1$$

$$\int_{-\infty}^{\hat{z}_{\text{MAE}}} f(z|X) dz = \frac{1}{2}, \quad \text{Conditional Median}$$

Maximum A Posteriory Probability (MAP):

$$C(\hat{Z}, Z) = \begin{cases} 1 & \text{when } \|\hat{Z} - Z\| > \delta \\ 0 & \text{when } \|\hat{Z} - Z\| \leq \delta \end{cases} \quad \delta \rightarrow 0$$

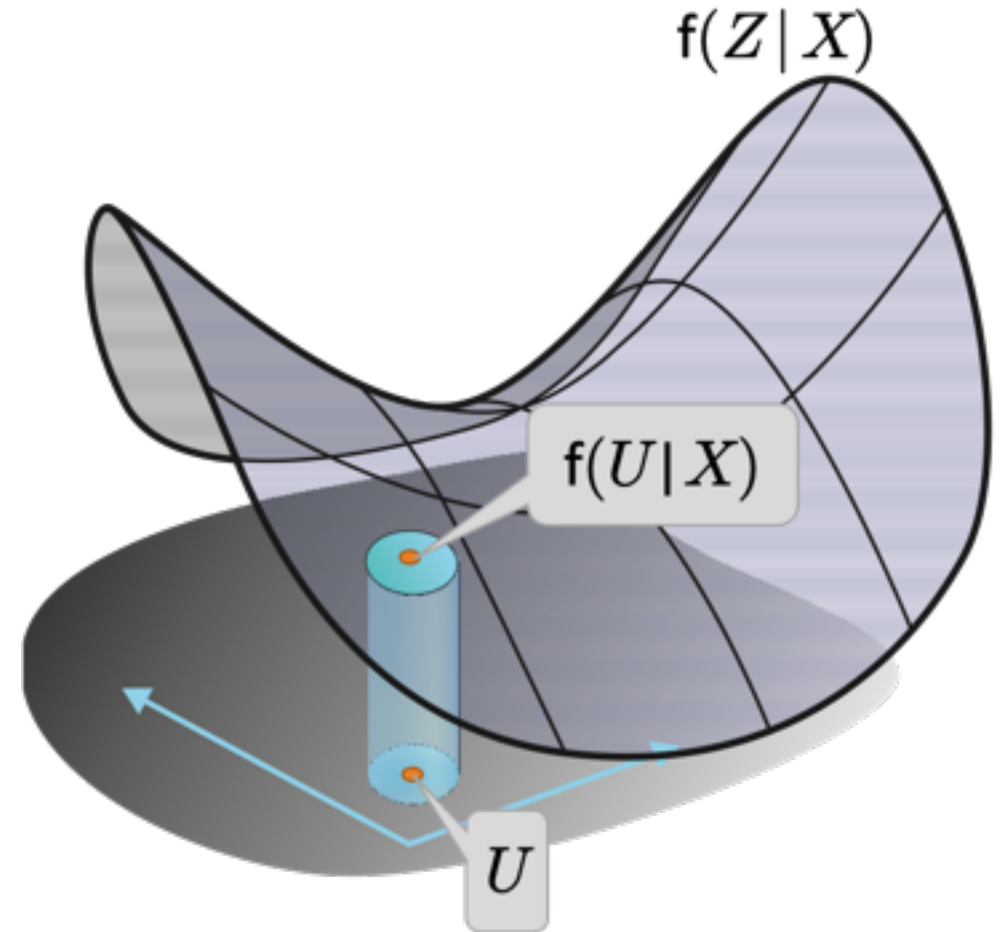
$$G(U, X) = \int C(U, Z) f(Z|X) dZ$$

$$= \int_{\|U-Z\| > \delta} f(Z|X) dZ$$

$$= 1 - \int_{\|U-Z\| \leq \delta} f(Z|X) dZ \approx 1 - \text{Ball}(\delta) f(U|X)$$

$$\min_U G(U, X) \approx 1 - \text{Ball}(\delta) \max_U f(U|X)$$

$$\hat{Z}_{\text{MAP}}(X) = \arg \max_Z f(Z|X)$$



Non-Bayesian Estimation

Measurements X and desired Z assumed to be “weakly” related through joint probability density $f(X, Z)$ considered **known**

From $f(X, Z)$ computed posterior probability density

$$f(Z|X) = \frac{f(X, Z)}{g(X)} = \frac{f(X, Z)}{\int f(X, Z) dZ}$$

In many applications no access to Z translates in no availability of $f(X, Z)$ and $f(Z|X)$

It is possible in some applications to know the conditional density $f(X|Z)$



$$X = Z + W$$

Conditional density $f(X|Z)$ requires only $g_w(W)$

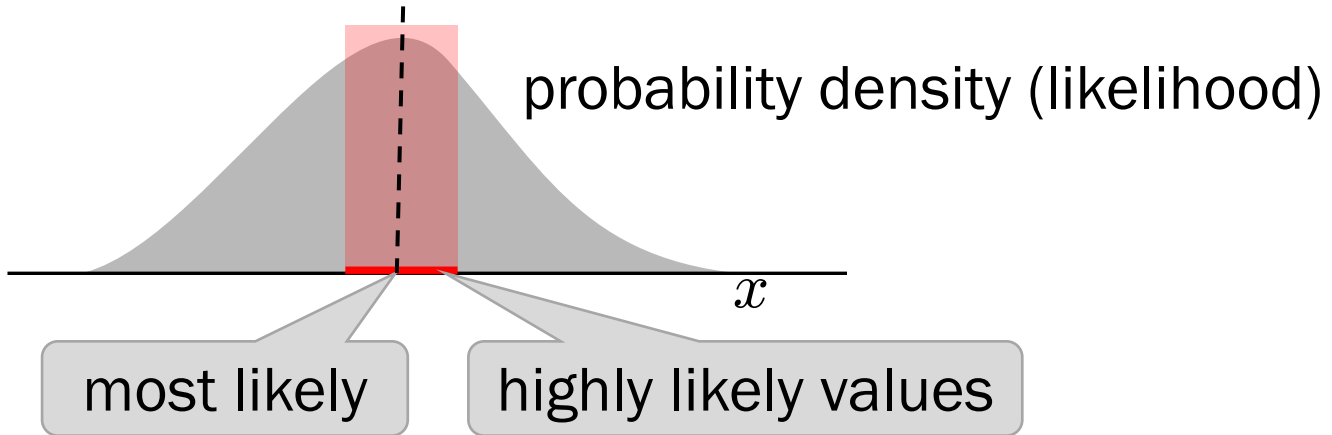
$$f(X|Z) = g_w(X - Z)$$

Joint density $f(X, Z)$ requires $g_w(W)$ and $h(Z)$

$$f(X, Z) = f(X|Z)h(Z) = g_w(X - Z)h(Z)$$

Knowing noise density $g_w(W)$ possible. Knowing $h(Z)$ difficult !

Maximum Likelihood Estimator (MLE):



Density $f(X|Z)$ is likelihood of X (given Z).

For measurements X what is Z that makes X **most likely** ?

We must solve $\max_Z f(X|Z)$

$$\hat{Z}_{\text{MLE}} = \arg \max_Z f(X|Z)$$

$$\hat{Z}_{\text{MAP}} = \arg \max_Z f(Z|X) = \arg \max_Z \frac{f(X, Z)}{g(X)} = \arg \max_Z f(X, Z)$$

$$= \arg \max_Z f(X|Z)h(Z)$$

if $h(Z)$ constant (degenerate uniform)

$$\hat{Z}_{\text{MLE}} = \arg \max_Z f(X|Z)$$

Optimality of MLE

Unbiased Estimator

$$\mathbb{E}_X [\hat{Z}(X)|Z] = Z$$

Interested in **Error covariance matrix**: $\mathbb{E}_X [(\hat{Z}(X) - Z)(\hat{Z}(X) - Z)^\top | Z]$

Cramer-Rao Lower Bound (CRLB)

From unbiased property

$$0 = \mathbb{E}_X \left[\hat{Z}(X) - Z | Z \right] = \int \left(\hat{Z}(X) - Z \right) f(X|Z) dX$$

Derivative (Jacobian) with respect to Z

$$0 = \int \left\{ -I f(X|Z) + \left(\hat{Z}(X) - Z \right) \left(\nabla_Z f(X|Z) \right)^\top \right\} dX$$

$$\begin{aligned} I &= \int \left\{ \left(\hat{Z}(X) - Z \right) \left(\frac{\nabla_Z f(X|Z)}{f(X|Z)} \right)^\top \right\} f(X|Z) dX \\ &= \mathbb{E}_X \left[\left(\hat{Z}(X) - Z \right) \left(\frac{\nabla_Z f(X|Z)}{f(X|Z)} \right)^\top \middle| Z \right] \end{aligned}$$

Define

$$\mathcal{E}(X) = \hat{Z}(X) - Z, \quad \Delta(X) = \frac{\nabla_Z f(X|Z)}{f(X|Z)}$$

Compute covariance matrix

$$\begin{aligned} \mathbb{E}_X \left[\begin{bmatrix} \mathcal{E}(X) \\ \Delta(X) \end{bmatrix} \begin{bmatrix} \mathcal{E}^\top(X) & \Delta^\top(X) \end{bmatrix} \middle| Z \right] &= \begin{bmatrix} \mathbb{E}_X[\mathcal{E}(X)\mathcal{E}^\top(X)|Z] & \mathbb{E}_X[\mathcal{E}(X)\Delta^\top(X)|Z] \\ \mathbb{E}_X[\Delta(X)\mathcal{E}^\top(X)|Z] & \mathbb{E}_X[\Delta(X)\Delta^\top(X)|Z] \end{bmatrix} \\ &= \begin{bmatrix} \mathbb{E}_X[\mathcal{E}(X)\mathcal{E}^\top(X)|Z] & I \\ I & \mathbb{E}_X[\Delta(X)\Delta^\top(X)|Z] \end{bmatrix} \end{aligned}$$

$$\mathbb{E}_X[\mathcal{E}(X)\mathcal{E}^\top(X)|Z] \geq (\text{FI})^{-1}$$

Fisher Information matrix

$$\text{FI} = \mathbb{E}_X[\Delta(X)\Delta^\top(X)|Z] = \mathbb{E}_X \left[\left(\frac{\nabla_Z f(X|Z)}{f(X|Z)} \right) \left(\frac{\nabla_Z f(X|Z)}{f(X|Z)} \right)^\top \middle| Z \right]$$

Estimation error covariance matrix of **any** unbiased estimator is lower bounded by inverse of Fisher Information matrix

$$\text{CRLB} = (\text{FI})^{-1}$$

Estimation error power of any estimator cannot go below a certain level

Call n the size of measurement data X

Theorem: Under general conditions we have

$$\lim_{n \rightarrow \infty} \mathbb{E}_X \left[(\hat{Z}_{\text{MLE}}(X) - Z) (\hat{Z}_{\text{MLE}}(X) - Z)^{\text{T}} | Z \right] \times \text{FI} = I$$

Asymptotic optimality (for large data size)

Summary

Minimum Mean Square Error: $\hat{Z}_{\text{MMSE}}(X) = \mathbb{E}[Z|X]$

Minimum Mean Absolute Error: $\hat{Z}_{\text{MMAE}}(X) = \arg \left\{ u : \int_{-\infty}^u f(z|X) dz = \frac{1}{2} \right\}$

Maximum A posteriori Probability: $\hat{Z}_{\text{MAP}}(X) = \arg \max_Z f(Z|X)$

Optimum for any measurement datasize.

Maximum Likelihood Estimator: $\hat{Z}_{\text{MLE}}(X) = \arg \max_Z f(X|Z)$

Asymptotically optimum, “large” measurement datasize.

Special Case

Let $Z = \{Z_1, Z_2\}$ and there is prior for Z_1 but not for Z_2

Treat non-existing priors as degenerate uniforms starting with MAP estimator

$$\begin{aligned}\arg \max_{Z_1, Z_2} f(Z_1, Z_2 | X) &= \arg \max_{Z_1, Z_2} f(X, Z_1, Z_2) \\ &= \arg \max_{Z_1, Z_2} f(X | Z_1, Z_2) h(Z_1, Z_2) \\ &= \arg \max_{Z_1, Z_2} f(X | Z_1, Z_2) h_1(Z_1 | Z_2) h_2(Z_2)\end{aligned}$$

$$h_2(Z_2) \text{ degenerate uniform} = \arg \max_{Z_1, Z_2} f(X | Z_1, Z_2) h_1(Z_1 | Z_2)$$

If Z_1 and Z_2 independent and interested only in estimating Z_1

$$\hat{Z}_1 = \arg \max_{Z_1} \left\{ \left(\max_{Z_2} f(X | Z_1, Z_2) \right) h_1(Z_1) \right\}$$