



Machine Learning Methods for Statistical Decision Making

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Two - Part Presentation

- Decision Making (Hypothesis Testing - Detection)
- Parameter Estimation

Goal: Consider the two well known problems under a pure data-driven setup

Decision Making: Outline

- Hypothesis Testing
 - Mathematical Formulation
 - Data Driven Approach
 - The Consistency Property
- Designing Consistent Test
 - Optimization Problems with Consistent Solutions
 - Data Driven Implementation
- Detection in Time Series
 - I.i.d. processes
 - Markov processes

Hypothesis Testing

Mathematical Formulation

For a random vector X we assume the following two hypotheses

$$H_0 : X \sim f_0(X), \mathbb{P}(H_0)$$

$$H_1 : X \sim f_1(X), \mathbb{P}(H_1)$$

For every X need to decide if it comes from H_0 or H_1

Decide using a *Decision Function* $D(X) \in \{0, 1\}$

Would like to **optimize** $D(X)$

Plethora of applications in diverse scientific fields!!!

Bayesian Approach

Minimize decision error probability

$$\min_D \left\{ \mathbb{P}(D = 1 | H_0) \mathbb{P}(H_0) + \mathbb{P}(D = 0 | H_1) \mathbb{P}(H_1) \right\}$$

$$\frac{f_1(X)}{f_0(X)} \underset{H_0}{\overset{H_1}{\geq}} \frac{\mathbb{P}(H_0)}{\mathbb{P}(H_1)} \equiv \frac{f_1(X) \mathbb{P}(H_1)}{f_0(X) \mathbb{P}(H_0)} \underset{H_0}{\overset{H_1}{\geq}} 1$$

For $\omega(r)$ strictly increasing

$$r(X) \underset{H_0}{\overset{H_1}{\geq}} 1 \equiv \omega(r(X)) \underset{H_0}{\overset{H_1}{\geq}} \omega(1), \quad r(X) = \frac{f_1(X) \mathbb{P}(H_1)}{f_0(X) \mathbb{P}(H_0)}$$

Neyman-Pearson Approach

$$H_0 : X \sim f_0(X), \mathbb{P}(H_0)$$

$$H_1 : X \sim f_1(X), \mathbb{P}(H_1)$$

Maximize detection probability $\mathbb{P}(D = 1|H_1)$

subject to false alarm probability constraint $\mathbb{P}(D = 1|H_0) \leq \alpha$

$$\frac{f_1(X)}{f_0(X)} \underset{H_0}{\overset{H_1}{\geq}} \lambda, \quad \mathbb{P}\left(\frac{f_1(X)}{f_0(X)} \geq \lambda \mid H_0\right) = \alpha$$

For $\omega(r)$ strictly increasing

$$\omega(r(X)) \underset{H_0}{\overset{H_1}{\geq}} \eta, \quad \mathbb{P}\left(\omega(r(X)) \geq \eta \mid H_0\right) = \alpha,$$

$$r(X) = \frac{f_1(X)}{f_0(X)}$$

Data Driven Approach

$$H_0 : X \sim \cancel{f_0(X)}, \cancel{\mathbb{P}(H_0)}$$

$$H_1 : X \sim \cancel{f_1(X)}, \cancel{\mathbb{P}(H_1)}$$

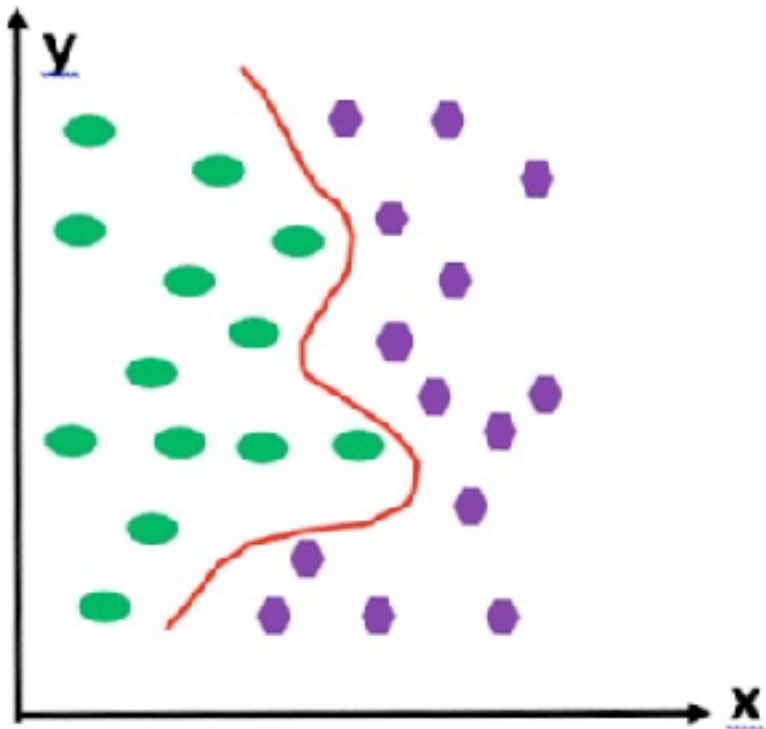
$$X_1^0 \ X_2^0 \ \dots \ X_{n_0}^0$$

Sampled from f_0

$$X_1^1 \ X_2^1 \ \dots \ X_{n_1}^1$$

Sampled from f_1

$$\mathbb{P}(H_i) \approx \frac{n_i}{n_0 + n_1}$$



Design **border** to separate the two datasets

What is the best border ?

$$\text{All } X : \frac{f_1(X)\mathbb{P}(H_1)}{f_0(X)\mathbb{P}(H_0)} = 1$$

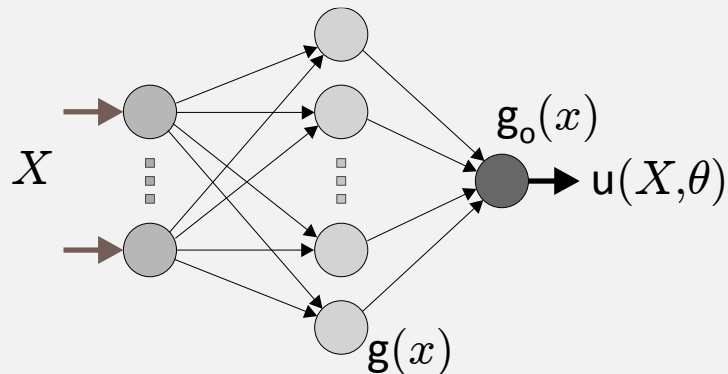
Instead of a “border”, design a decision like function $v(X)$

$$v(X) = \begin{cases} -1 & \text{when } X \text{ from } H_0 \\ 1 & \text{when } X \text{ from } H_1. \end{cases}$$

Designing a function $v(X)$ when information is in the form of data is challenging. We replace $v(X)$ with a neural network $u(X, \theta)$ and select parameters θ

Cybenko 1989 (universal approximation)

For sufficiently large neural network $u(X, \theta)$ we can find suitable parameters θ such that we can approximate arbitrarily close any function $v(X)$



$$|v(X) - u(X, \theta)| \leq \epsilon$$

Select a neural network configuration $u(X, \theta)$ and optimize network parameters θ by defining the distance

$$J(\theta) = \frac{1}{n_0 + n_1} \left\{ \sum_{i=1}^{n_0} \left(-1 - u(X_i^0, \theta) \right)^2 + \sum_{j=1}^{n_1} \left(1 - u(X_j^1, \theta) \right)^2 \right\}$$

and solving the optimization problem

$$\min_{\theta} J(\theta) \Rightarrow \theta_0 \Rightarrow u(X, \theta_0)$$

We use the resulting function $u(X, \theta_0)$ to make a decision for any new data X as follows:

$$u(X, \theta_0) \begin{matrix} > \\ \equiv \\ < \end{matrix} \begin{matrix} H_1 \\ \mathbf{0} \\ H_0 \end{matrix}$$

Works “well”!! Why??

Understanding using Asymptotic Analysis

$$n_0, n_1 \rightarrow \infty, \quad u(X, \theta) \rightarrow v(X)$$

$$J(\theta) = \frac{n_0}{n_0 + n_1} \frac{1}{n_0} \sum_{i=1}^{n_0} \left(1 + u(X_i^0, \theta)\right)^2 + \frac{n_1}{n_0 + n_1} \frac{1}{n_1} \sum_{j=1}^{n_1} \left(1 - u(X_j^1, \theta)\right)^2$$

$$J(v) = \mathbb{P}(H_0) \mathbb{E}_0 \left[\left(1 + v(X)\right)^2 \right] + \mathbb{P}(H_1) \mathbb{E}_1 \left[\left(1 - v(X)\right)^2 \right]$$

$$\min_{\theta} J(\theta) \rightarrow \min_v J(v)$$

$$\theta_o \Rightarrow u(X, \theta_o) \approx v_o(X)$$

$$\mathbb{E}_1 \left[\left(1 - v(X)\right)^2 \right] = \mathbb{E}_0 \left[\left(1 - v(X)\right)^2 \frac{f_1(X)}{f_0(X)} \right]$$

$$J(v) = \mathbb{P}(H_0) \mathbb{E}_0 \left[\left(1 + v(X)\right)^2 + r(X) \left(1 - v(X)\right)^2 \right] \quad r(X) = \frac{f_1(X) \mathbb{P}(H_1)}{f_0(X) \mathbb{P}(H_0)}$$

minimize for each fixed X

$$v_o(X) = \frac{r(X) - 1}{r(X) + 1} = \omega(r(X)), \quad \text{where } \omega(r) = \frac{r - 1}{r + 1} \text{ strictly increasing}$$

$$\text{Test equivalent to Bayes: } v_o(X) = \omega(r(X)) \underset{H_0}{\overset{H_1}{\geq}} \omega(1) = 0 \Rightarrow u(X, \theta_o) \underset{H_0}{\overset{H_1}{\geq}} 0$$

Equivalence in the limit

Consistency (with respect to the Bayes test)

$$\min_{\theta} J(\theta) = \min_{\theta} \left\{ \sum_{i=1}^{n_0} \left(1 + u(X_i^0, \theta)\right)^2 + \sum_{j=1}^{n_1} \left(1 - u(X_j^1, \theta)\right)^2 \right\}$$

$$\Rightarrow \theta_0 \Rightarrow u(X, \theta_0)$$

$$\min_{\nu} J(\nu) = \min_{\nu} \left\{ \mathbb{P}(H_0) \mathbb{E}_0 \left[\left(1 + \nu(X)\right)^2 \right] + \mathbb{P}(H_1) \mathbb{E}_1 \left[\left(1 - \nu(X)\right)^2 \right] \right\}$$

$$\Rightarrow \nu_0(X) = \omega(r(X))$$

Expect: $u(X, \theta_0) \approx \omega(r(X))$

Optimum Test: $\omega(r(X)) \underset{H_0}{\overset{H_1}{\gtrless}} \omega(1)$, Close to Optimum: $u(X, \theta_0) \underset{H_0}{\overset{H_1}{\gtrless}} \omega(1)$

Develop data driven methods for estimation of $\omega(r(X))$ for other $\omega(r)$

Consistent tests **eventually** prevail over inconsistent tests

Designing Consistent Test

$$r(X) \underset{H_0}{\overset{H_1}{\gtrless}} 1 \equiv \omega_1(r(X)) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_1(1) \equiv \omega_2(r(X)) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_2(1)$$

$$u(X, \theta_0) \underset{H_0}{\overset{H_1}{\gtrless}} 1 \not\equiv u_1(X, \theta_1) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_1(1) \not\equiv u_2(X, \theta_2) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_2(1)$$

For function $\omega(r)$ can we define cost

$$J(v) = \mathbb{P}(H_0) \mathbb{E}_0 \left[\left(\frac{v(X) - \mathbb{E}_0[v(X)]}{\mathbb{E}_0[v(X)]} \right)^2 \right] + \mathbb{P}(H_1) \mathbb{E}_1 \left[\left(\frac{v(X) - \mathbb{E}_1[v(X)]}{\mathbb{E}_1[v(X)]} \right)^2 \right]$$

so that $\min_v J(v) \Rightarrow v_0(X) = \omega(r(X))$?

THEOREM: Select **strictly increasing** function $\omega(r)$ and **strictly negative** function $\rho(z)$. Define

$$\psi'(z) = \rho(z), \quad \phi'(z) = -\omega^{-1}(z)\rho(z)$$

then the solution of the optimization problem

$$\min_{\mathbf{v}} J(\mathbf{v}) = \min_{\mathbf{v}} \left\{ \mathbb{P}(\mathbf{H}_0)\mathbb{E}_0 [\phi(\mathbf{v}(X))] + \mathbb{P}(\mathbf{H}_1)\mathbb{E}_1 [\psi(\mathbf{v}(X))] \right\}$$

satisfies $\mathbf{v}_0(X) = \arg \min_{\mathbf{v}} J(\mathbf{v}) = \omega(r(X))$

Same optimal solution $\omega(r(X))$ for all functions $\rho(z) < 0$

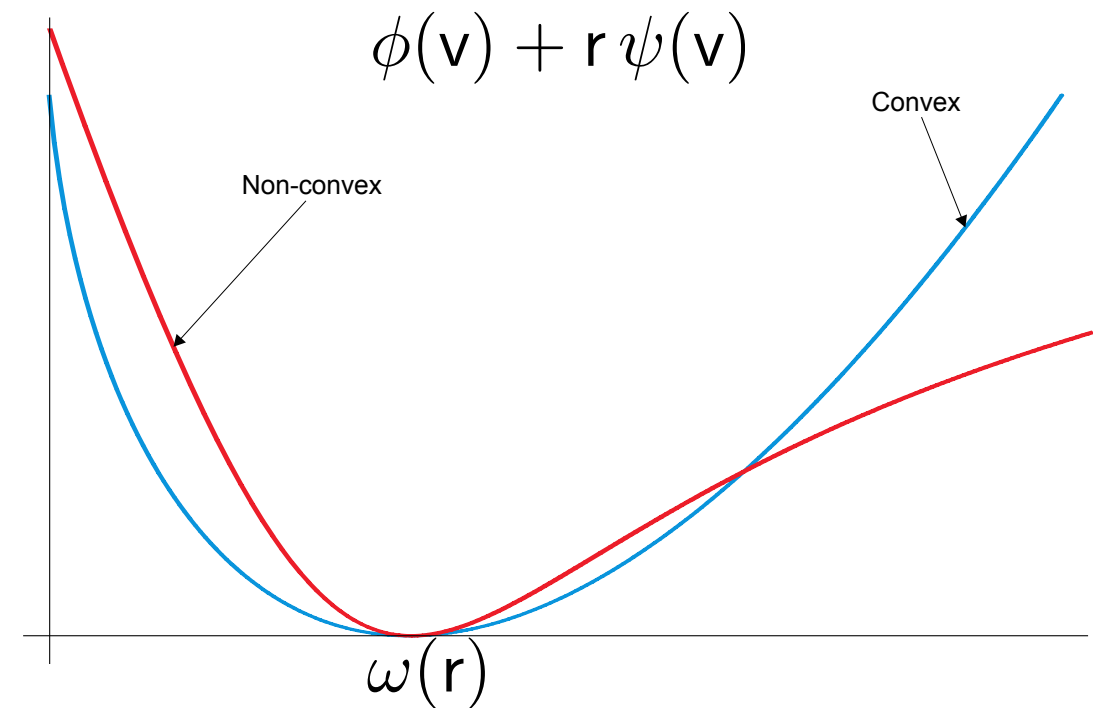
Proof (highlights): Apply change of measures in the second part

$$\begin{aligned} J(\mathbf{v}) &= \mathbb{P}(\mathbf{H}_0)\mathbb{E}_0 [\phi(\mathbf{v}(X))] + \mathbb{P}(\mathbf{H}_1)\mathbb{E}_1 [\psi(\mathbf{v}(X))] \\ &= \mathbb{P}(\mathbf{H}_0)\mathbb{E}_0 [\phi(\mathbf{v}(X)) + r(X)\psi(\mathbf{v}(X))] \end{aligned}$$

therefore

$$\begin{aligned} \min_{\mathbf{v}} J(\mathbf{v}) \\ &= \min_{\mathbf{v}} \mathbb{P}(\mathbf{H}_0)\mathbb{E}_0 [\phi(\mathbf{v}(X)) + r(X)\psi(\mathbf{v}(X))] \end{aligned}$$

minimize for each fixed X



Examples of functions

A: $\omega(r) = r \in \mathbb{R}_+$ (likelihood ratio)

$$\rho(z) = -1, z \geq 0 \Rightarrow \phi(z) = \frac{z^2}{2}, \psi(z) = -z$$

Mean
Square

B: $\omega(r) = \log(r) \in \mathbb{R}$ (log-likelihood ratio)

$$\rho(z) = -e^{-0.5z} \Rightarrow \phi(z) = 2e^{0.5z}, \psi(z) = 2e^{-0.5z}$$

Exponential

C: $\omega(r) = \frac{r}{r+1} \in [0, 1]$ (posterior probability)

$$\rho(z) = -\frac{1}{z}, z \in [0, 1] \Rightarrow \phi(z) = -\log(1-z), \psi(z) = -\log(z)$$

Cross
Entropy

Data Driven Implementation

$$J(\mathbf{v}) = \mathbb{P}(\mathbf{H}_0)\mathbb{E}_0 [\phi(\mathbf{v}(X))] + \mathbb{P}(\mathbf{H}_1)\mathbb{E}_1 [\psi(\mathbf{v}(X))]$$

$$J(\theta) = \frac{1}{n_0 + n_1} \left\{ \sum_{i=1}^{n_0} \phi(\mathbf{u}(X_i^0, \theta)) + \sum_{j=1}^{n_1} \psi(\mathbf{u}(X_j^1, \theta)) \right\}$$
$$\mathbf{u}(X, \theta_0) \approx \omega \left(\frac{f_1(X)\mathbb{P}(\mathbf{H}_1)}{f_0(X)\mathbb{P}(\mathbf{H}_0)} \right)$$

$$J(\theta) = \frac{1}{n_0} \sum_{i=1}^{n_0} \phi(\mathbf{u}(X_i^0, \theta)) + \frac{1}{n_1} \sum_{j=1}^{n_1} \psi(\mathbf{u}(X_j^1, \theta))$$
$$\mathbf{u}(X, \theta_0) \approx \omega \left(\frac{f_1(X)}{f_0(X)} \right)$$

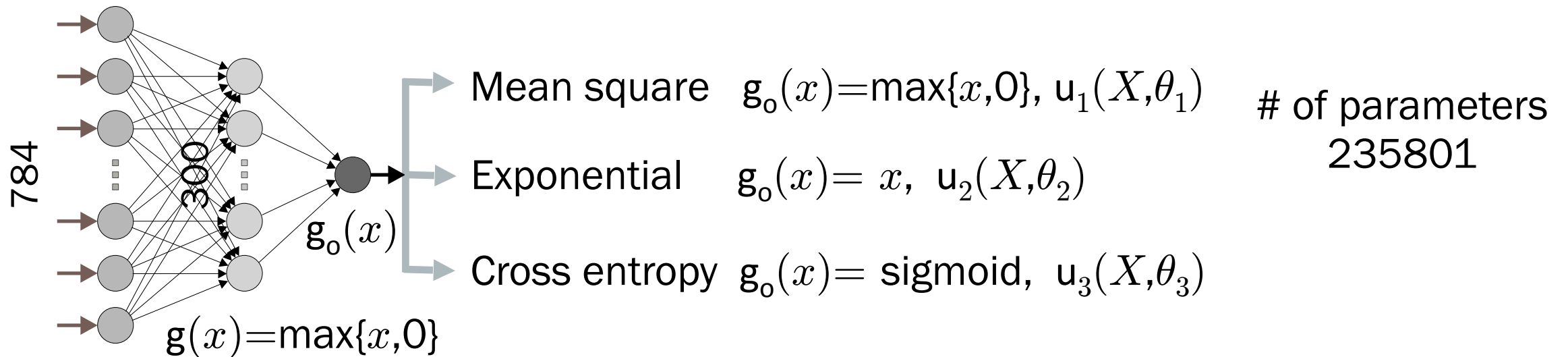
Example: Classification Problem

From dataset MNIST isolate handwritten numerals 4 and 9



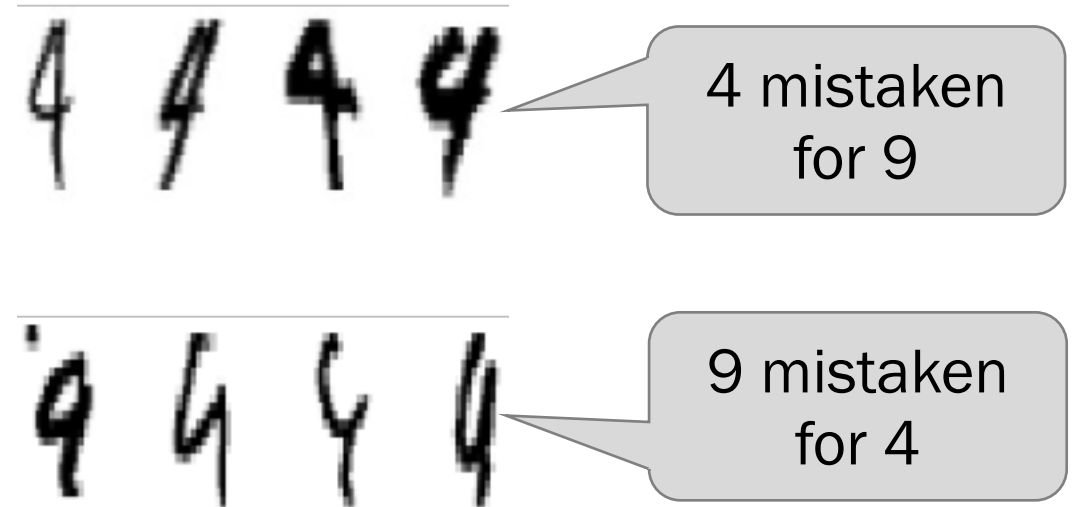
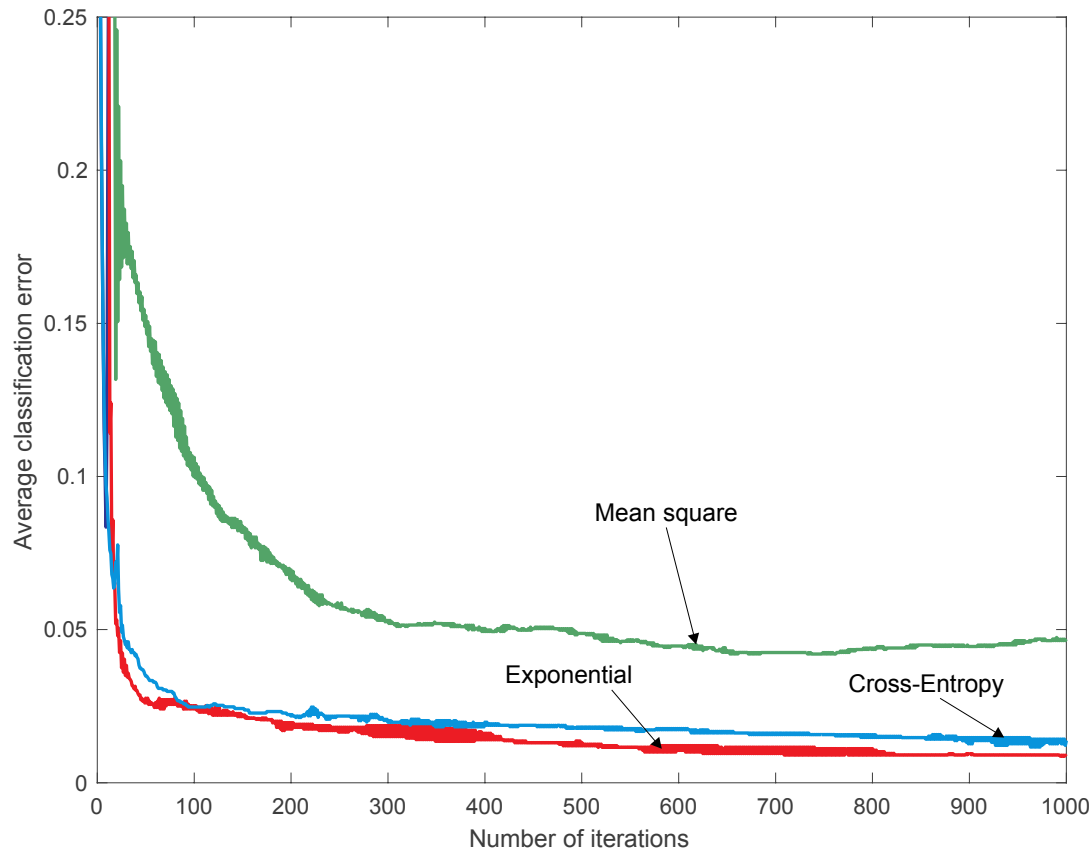
Gray scale images $28 \times 28 = 784$ pixels. Design classifier using training data. Examine performance using testing data.

Neural network $784 \times 300 \times 1$



Training set: 5500 “4” and 5500 “9”. Testing set: 982 “4” and 1009 “9”

$$u_1(X, \theta_1) \underset{H_0}{\overset{H_1}{\gg}} 1, \quad u_2(X, \theta_2) \underset{H_0}{\overset{H_1}{\gg}} 0, \quad u_3(X, \theta_3) \underset{H_0}{\overset{H_1}{\gg}} \frac{1}{2}$$



Detection in Time Series

More practically interesting case: Testing of time series $\{ X_1, X_2, \dots, X_n \}$

The **whole** set of measurements under H_0 or H_1

For testing we need likelihood ratio

$$\frac{f_1(X_n, \dots, X_1)}{f_0(X_n, \dots, X_1)} = \frac{f_1(X_n | X_{n-1}, \dots, X_1)}{f_0(X_n | X_{n-1}, \dots, X_1)} \frac{f_1(X_{n-1} | X_{n-2}, \dots, X_1)}{f_0(X_{n-1} | X_{n-2}, \dots, X_1)} \dots \frac{f_1(X_1)}{f_0(X_1)}$$

When i.i.d. under each hypothesis

$$\frac{f_1(X_n, \dots, X_1)}{f_0(X_n, \dots, X_1)} = \frac{f_1(X_n)}{f_0(X_n)} \dots \frac{f_1(X_1)}{f_0(X_1)}$$

Test to be used

$$\sum_{i=1}^n \log \left(\frac{f_1(X_i)}{f_0(X_i)} \right) \underset{H_0}{\overset{H_1}{\gtrless}} \eta$$

Interested in estimating $\omega(r(X)) = \omega\left(\frac{f_1(X)}{f_0(X)}\right)$

We are given training data: $\{X_1^0, \dots, X_{n_0}^0\}$ following H_0
 $\{X_1^1, \dots, X_{n_1}^1\}$ following H_1

For each $\omega(r)$ of interest, minimize corresponding $J(\theta)$

$$J(\theta) = \frac{1}{n_0} \sum_{i=1}^{n_0} \phi(u(X_i^0, \theta)) + \frac{1}{n_1} \sum_{j=1}^{n_1} \psi(u(X_j^1, \theta))$$

$$u(X, \theta_0) \approx \omega\left(\frac{f_1(X)}{f_0(X)}\right)$$

For $\omega(r) = r$, $u_1(X, \theta_1) \approx \frac{f_1(X)}{f_0(X)}$, use $\log(u_1(X, \theta_1))$ (Mean Square)

For $\omega(r) = \log(r)$, $u_2(X, \theta_2) \approx \log\left(\frac{f_1(X)}{f_0(X)}\right)$, use $u_2(X, \theta_2)$ (Exponential)

For $\omega(r) = \frac{r}{r+1}$, $u_3(X, \theta_3) \approx \frac{\frac{f_1(X)}{f_0(X)}}{\frac{f_1(X)}{f_0(X)} + 1}$, use $\log\left(\frac{u_3(X, \theta_3)}{1 - u_3(X, \theta_3)}\right)$
(Cross Entropy)

Example: Testing i.i.d. sequences

Assume $\{X_i\}$ are vectors of length 10

We would like to test 20 consecutive samples $\{X_1, \dots, X_{20}\}$

$$H_0: f_0 \sim \mathcal{N}(0, I)$$

$$H_1: f_1 \sim \mathcal{N}\left(\frac{1}{\sqrt{10}}[1 \cdots 1], 1.2I\right)$$

Neural Network $10 \times 20 \times 1$

Training data $n_0 = n_1 = 100$

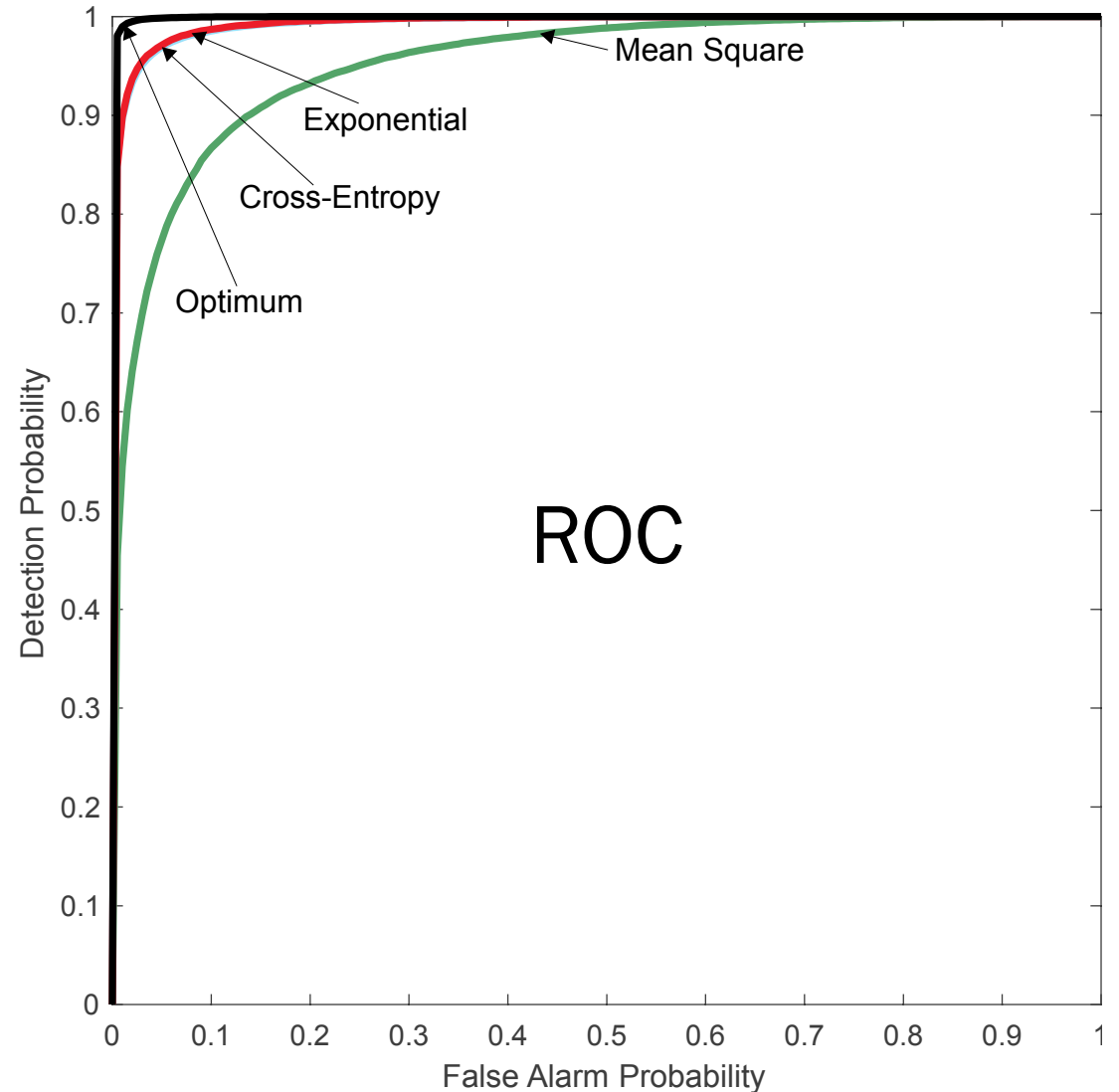
Produce $u_1(X, \theta_1), u_2(X, \theta_2), u_3(X, \theta_3)$

We will employ $\log(u_1(X, \theta_1)), u_2(X, \theta_2), \log\left(\frac{u_3(X, \theta_3)}{1 - u_3(X, \theta_3)}\right)$

Testing for $n = 20$ samples

$$\left. \begin{array}{l} \sum_{i=1}^{20} \log(u_1(X_i, \theta_1)) \\ \sum_{i=1}^{20} u_2(X_i, \theta_2) \\ \sum_{i=1}^{20} \log\left(\frac{u_3(X_i, \theta_3)}{1-u_3(X_i, \theta_3)}\right) \end{array} \right\} \begin{array}{l} H_1 \\ \eta \\ H_0 \end{array}$$

We generate 100000×20 realizations from f_0 and from f_1



Testing Markovian processes

Consider Markovian processes with “memory” m

$$\frac{f_1(X_n | X_{n-1}, \dots, X_1)}{f_0(X_n | X_{n-1}, \dots, X_1)} = \frac{f_1(X_n | X_{n-1}, \dots, X_{n-m})}{f_0(X_n | X_{n-1}, \dots, X_{n-m})}$$

$$\frac{f_1(X_n, \dots, X_1)}{f_0(X_n, \dots, X_1)} = \frac{f_1(X_n | X_{n-1}, \dots, X_{n-m})}{f_0(X_n | X_{n-1}, \dots, X_{n-m})} \dots \frac{f_1(X_{m+1} | X_m, \dots, X_1)}{f_0(X_{m+1} | X_m, \dots, X_1)} \times \frac{f_1(X_m, \dots, X_1)}{f_0(X_m, \dots, X_1)}$$

Can we estimate likelihood ratio of conditional densities?

- a) Through data dynamics (classical)
- b) Directly (proposed)

Classical Approach

Most common model, Autoregressive

$$X_t = A_1^i X_{t-1} + \dots + A_m^i X_{t-m} + W_t, \quad i = 0, 1$$

$$X_t = G_i(X_{t-1}, \dots, X_{t-m}, \theta^i) + W_t$$

Use training data $\{X_1^0, \dots, X_{n_0}^0\}$ and $\{X_1^1, \dots, X_{n_1}^1\}$ to solve

$$\min_{\theta^i} \sum_{t=1}^{n_i} \left(X_t^i - G_i(X_{t-1}^i, \dots, X_{t-m}^i, \theta^i) \right)^2 \Rightarrow \theta_0^i$$

$$W_t^i = X_t^i - G_i(X_{t-1}^i, \dots, X_{t-m}^i, \theta_0^i), \quad \Sigma_i = \frac{1}{n_i} \sum_{t=1}^{n_i} W_t^i (W_t^i)^\top$$

Assume $\{W_t^i\}$ i.i.d. Gaussian $\mathcal{N}(0, \Sigma_i)$

X_t given $\{X_{t-1}, \dots, X_{t-m}\}$ under hypothesis H_i

Gaussian with mean $G_i(X_{t-1}, \dots, X_{t-m}, \theta_0^i)$ and covariance Σ_i

To test $\{X_1, \dots, X_n\}$

$$W_t^i = X_t - G_i(X_{t-1}, \dots, X_{t-m}, \theta_0^i)$$

$$\frac{f_1(X_t | X_{t-1}, \dots, X_{t-m})}{f_0(X_t | X_{t-1}, \dots, X_{t-m})} = \frac{e^{-\frac{1}{2}(W_t^1)^\top \Sigma_1^{-1} W_t^1}}{e^{-\frac{1}{2}(W_t^0)^\top \Sigma_0^{-1} W_t^0}} \sqrt{\frac{|\Sigma_0|}{|\Sigma_1|}}$$

Not purely data driven

Gaussian assumption arbitrary, not necessarily suitable for all data!

Proposed Approach

$$\log \left(\frac{f_1(X_t | X_{t-1}, \dots, X_{t-m})}{f_0(X_t | X_{t-1}, \dots, X_{t-m})} \right) = \log \left(\frac{\frac{f_1(X_t, X_{t-1}, \dots, X_{t-m})}{f_1(X_{t-1}, \dots, X_{t-m})}}{\frac{f_0(X_t, X_{t-1}, \dots, X_{t-m})}{f_0(X_{t-1}, \dots, X_{t-m})}} \right) =$$

$$\log \left(\frac{f_1(X_t, X_{t-1}, \dots, X_{t-m})}{f_0(X_t, X_{t-1}, \dots, X_{t-m})} \right)$$

$$\mathbf{u}_{m+1}(X_t, \dots, X_{t-m}, \theta_{m+1})$$

$$- \log \left(\frac{f_1(X_{t-1}, \dots, X_{t-m})}{f_0(X_{t-1}, \dots, X_{t-m})} \right)$$

$$\mathbf{u}_m(X_{t-1}, \dots, X_{t-m}, \theta_m)$$

$$\log \left(\frac{f_1(X_t | X_{t-1}, \dots, X_{t-m})}{f_0(X_t | X_{t-1}, \dots, X_{t-m})} \right) \approx$$

$$\mathbf{u}_{m+1}(X_t, \dots, X_{t-m}, \theta_{m+1}) - \mathbf{u}_m(X_{t-1}, \dots, X_{t-m}, \theta_m)$$

Example: Testing Markov sequences

Scalar observations $\{x_1, \dots, x_n\}$

$w_t \sim \mathcal{N}(0, 1)$, i.i.d.

$H_0 : x_t = w_t$

$H_1 : x_t = \text{sign}(x_{t-1})\sqrt{|x_{t-1}|} + w_t$

$u_2(x_t, x_{t-1}, \theta_2) : 2 \times 20 \times 1$

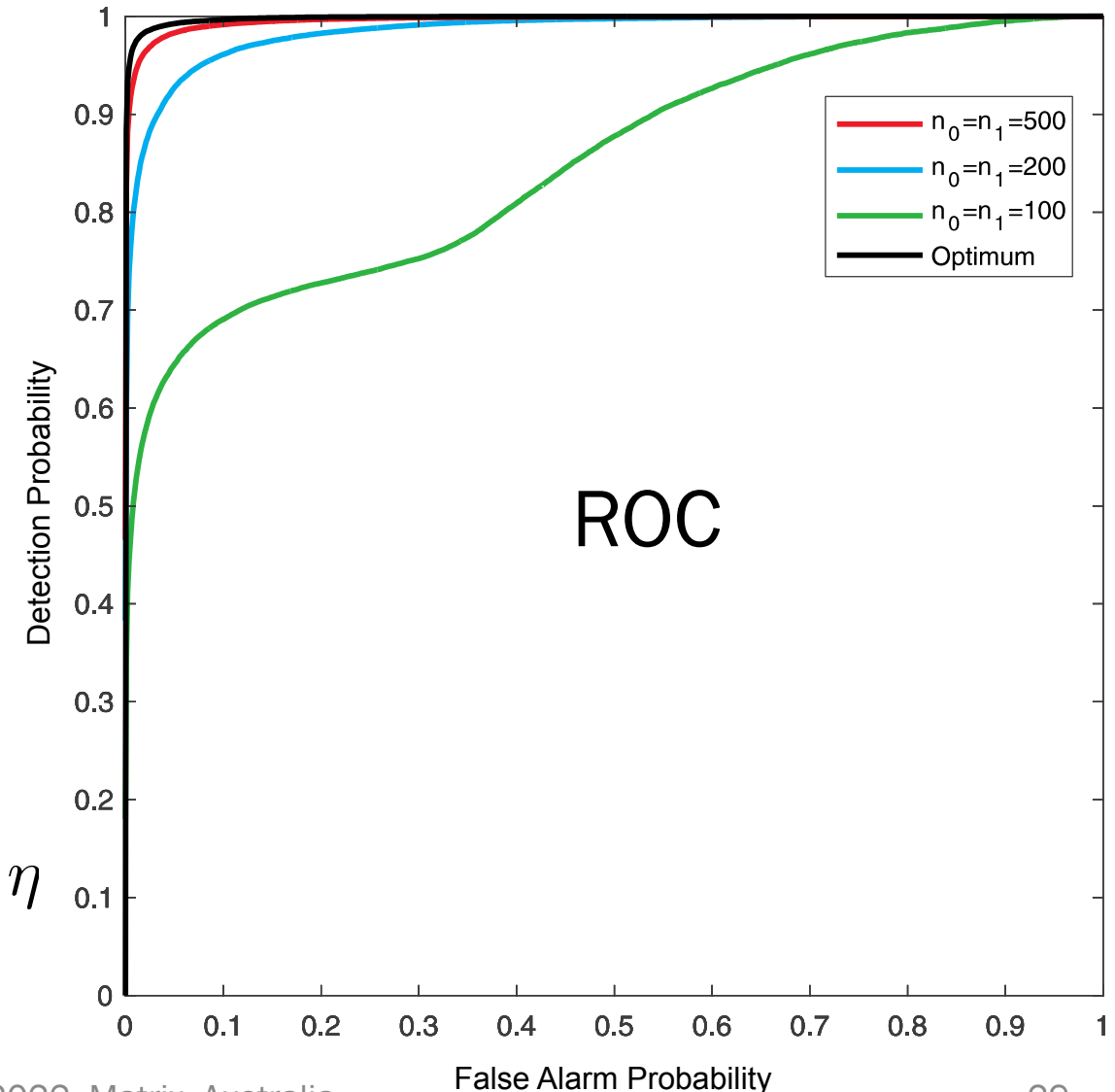
$u_1(x_t, \theta_1) : 1 \times 10 \times 1$ (Exponential)

Training data $n_0 = n_1 = 100, 200, 500$

Testing $n = 20$, i.e. $\{x_1, \dots, x_{20}\}$

$$\sum_{t=2}^{20} u_2(x_t, x_{t-1}, \theta_2) - \sum_{t=2}^{19} u_1(x_t, \theta_1) \underset{H_0}{\overset{H_1}{\gtrless}} \eta$$

100000 \times 20 samples from H_0 and H_1



Parameter Estimation: Outline

- Probability density vs Generative model
 - Inverse problems

Probability Density vs Generative Model

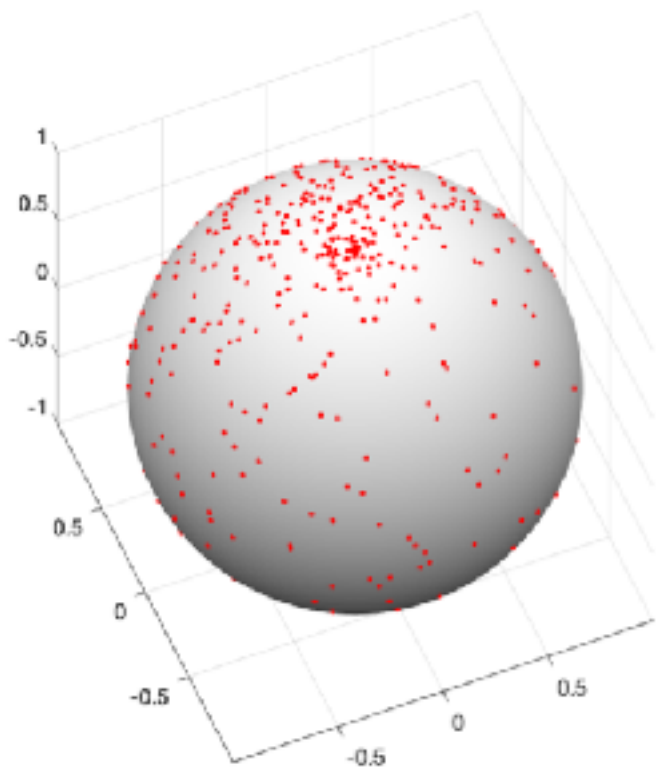
Points in N-D space can be random and lie on a lower dimensional surface (manifold)

Example red points on sphere (2-D in 3-D space)

Points are random with coordinates $[y_1, y_2, y_3]$ related through a **deterministic** equation

To lie on a sphere of radius r : $y_1^2 + y_2^2 + y_3^2 = r^2$

$$f(y_1, y_2, y_3) = \delta(y_1^2 + y_2^2 + y_3^2 - r^2) h(y_1, y_2)$$



Generative model: Data are representable as $Y = G(Z)$, $Z \sim h(Z)$. Many datasets satisfy

$$\dim(Y) \gg \dim(Z)$$

To design $G(Z)$ we assume existence of training set $\{Y_1, \dots, Y_n\}$

Approximate Generator with neural network $G(Z, \theta)$

Define second neural network the Discriminator $D(X, \vartheta)$

$$\min_{\theta} \max_{\vartheta} \left\{ \frac{1}{n} \sum_{t=1}^n \phi(D(Y_t, \vartheta)) + \frac{1}{m} \sum_{t=1}^m \psi(D(G(Z_t, \theta), \vartheta)) \right\}$$

$\Rightarrow \theta_o \Rightarrow G(Z, \theta_o)$

Generative
Adversarial
Network

Generator $G(Z, \theta_o)$ when applied to realizations of Z yields samples following closely the density of $\{Y_1, \dots, Y_n\}$

Example

HD-CelebA (30 000 high definition images 1024 X 1024 of celebrities)

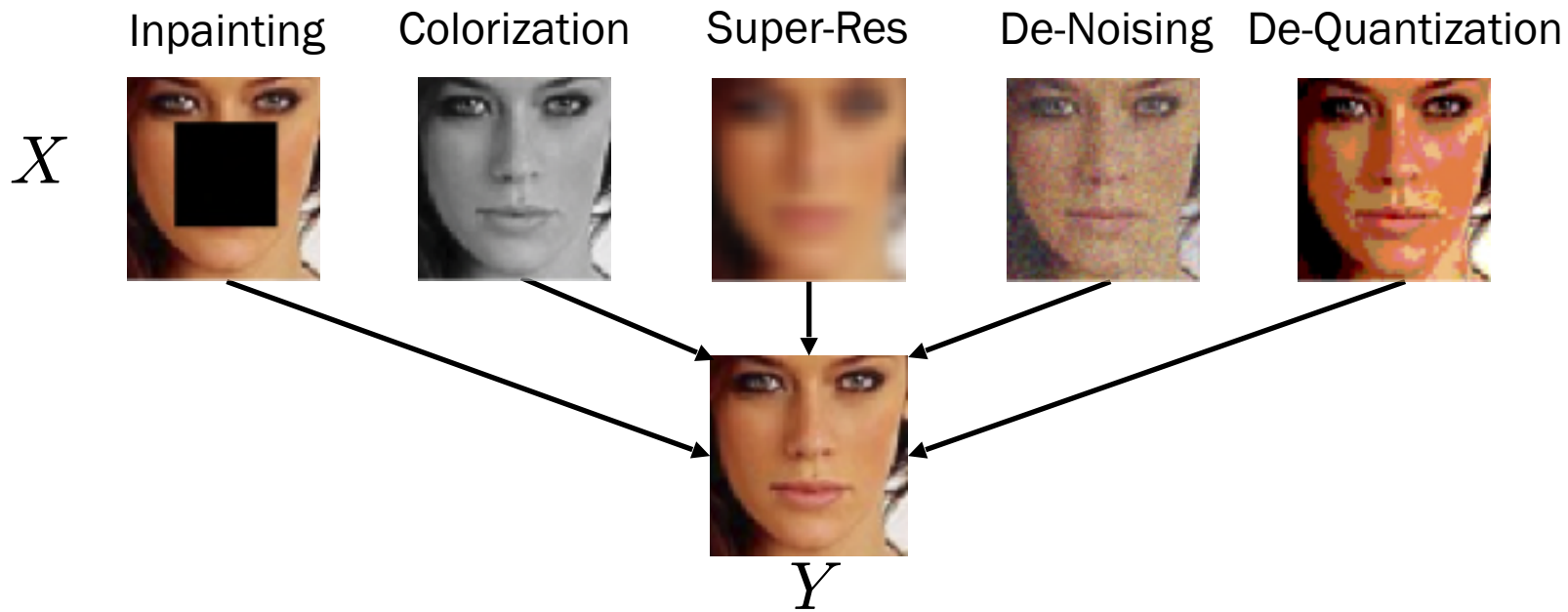


Extremely **hard to control convergence** of the adversarial problem

NVIDIA used progressive growing of GANs (4X4), (8X8),..., (1024X1024)

Design of $Y=G(Z,\theta)$ where Z is Gaussian vector of length 500





$$X = T(Y) + W$$

Instead of estimating Y from X , since $Y=G(Z)$, we first estimate Z and then recover Y from $Y=G(Z)$

Instead of estimating 3×10^6 variables from X , we only estimate 500 (vector Z)

Y : Follows generative model $G(Z)$

Z : Input is Gaussian with mean 0 and covariance identity

X : Measurement is vector of length N

We can estimate input Z by solving the optimization problem

$$\min_Z \left\{ \log \left(\|X - T(G(Z))\|^2 \right) + \frac{1}{N} \|Z\|^2 \right\}$$

$$\Rightarrow \hat{Z} \Rightarrow \hat{Y} = G(\hat{Z})$$

Optimization problem is an outcome of rigorous analysis based on Statistical estimation theory where probability densities are replaced by generative models

Blurring with 3 X 3 mask



Colorization (green channel)



De-Quantization

3 levels per RGB channel, 27 colors

