



# Data Driven Detection Methods

GEORGE V. MOUSTAKIDES  
UNIVERSITY OF PATRAS, GREECE

# Outline

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# Hypothesis Testing

## Mathematical Formulation

For a random vector  $X$  we assume the following two hypotheses

$$H_0 : X \sim f_0(X), \mathbb{P}(H_0)$$

$$H_1 : X \sim f_1(X), \mathbb{P}(H_1)$$

For every  $X$  need to decide if it comes from  $H_0$  or  $H_1$

Decide using a *Decision Function*  $D(X) \in \{0, 1\}$

Would like to **optimize**  $D(X)$

Plethora of applications in diverse scientific fields!!!

# Bayesian Approach

Minimize decision error probability

$$\min_D \left\{ \mathbb{P}(D = 1 | H_0) \mathbb{P}(H_0) + \mathbb{P}(D = 0 | H_1) \mathbb{P}(H_1) \right\}$$

$$\frac{f_1(X)}{f_0(X)} \underset{H_0}{\overset{H_1}{\geq}} \frac{\mathbb{P}(H_0)}{\mathbb{P}(H_1)} \equiv \frac{f_1(X) \mathbb{P}(H_1)}{f_0(X) \mathbb{P}(H_0)} \underset{H_0}{\overset{H_1}{\geq}} 1$$

For  $\omega(r)$  strictly increasing

$$r(X) \underset{H_0}{\overset{H_1}{\geq}} 1 \equiv \omega(r(X)) \underset{H_0}{\overset{H_1}{\geq}} \omega(1), \quad r(X) = \frac{f_1(X) \mathbb{P}(H_1)}{f_0(X) \mathbb{P}(H_0)}$$

## Neyman-Pearson Approach

$$\begin{aligned} H_0 : X &\sim f_0(X), \mathbb{P}(H_0) \\ H_1 : X &\sim f_1(X), \mathbb{P}(H_1) \end{aligned}$$

Maximize detection probability  $\mathbb{P}(D = 1|H_1)$

subject to false alarm probability constraint  $\mathbb{P}(D = 1|H_0) \leq \alpha$

$$\frac{f_1(X)}{f_0(X)} \underset{H_0}{\overset{H_1}{\geq}} \lambda, \quad \mathbb{P} \left( \frac{f_1(X)}{f_0(X)} \geq \lambda \middle| H_0 \right) = \alpha$$

For  $\omega(r)$  strictly increasing

$$\omega(r(X)) \underset{H_0}{\overset{H_1}{\geq}} \eta, \quad \mathbb{P} \left( \omega(r(X)) \geq \eta \middle| H_0 \right) = \alpha,$$

$$r(X) = \frac{f_1(X)}{f_0(X)}$$

# Data Driven Approach

$$H_0 : X \sim f_0(X), \mathbb{P}(H_0)$$

$$H_1 : X \sim f_1(X), \mathbb{P}(H_1)$$

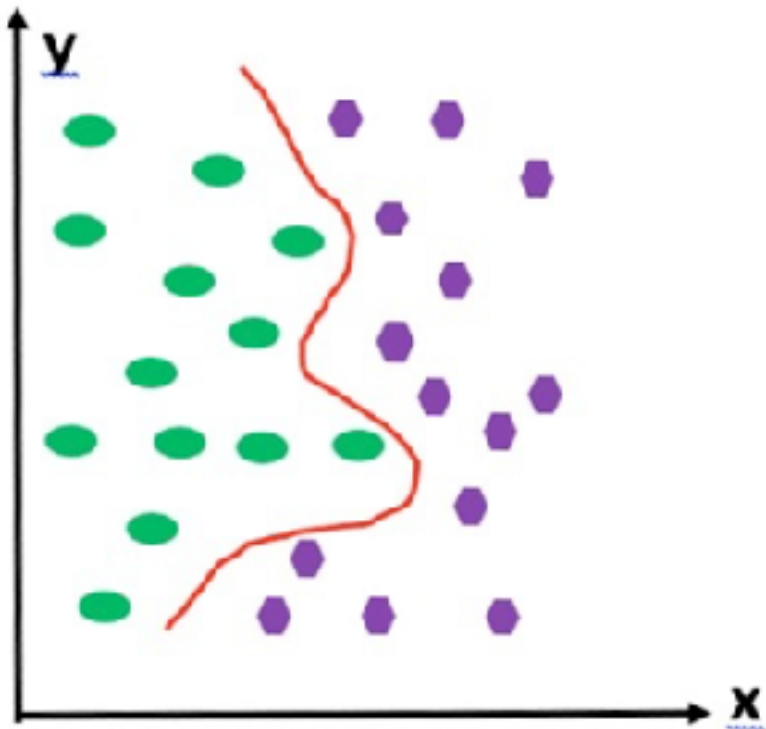
$$X_1^0 \ X_2^0 \ \dots \ X_{n_0}^0$$

$$X_1^1 \ X_2^1 \ \dots \ X_{n_1}^1$$

Sampled from  $f_0$

Sampled from  $f_1$

$$\mathbb{P}(H_i) \approx \frac{n_i}{n_0 + n_1}$$



Design **border** to separate the two datasets

What is the best border ?

$$\text{All } X : \frac{f_1(X)\mathbb{P}(H_1)}{f_0(X)\mathbb{P}(H_0)} = 1$$

Instead of a “border”, design a decision like function  $v(X)$

$$v(X) = \begin{cases} -1 & \text{when } X \text{ from } H_0 \\ 1 & \text{when } X \text{ from } H_1. \end{cases}$$

Use parametric family of functions  $u(X, \theta)$  and optimize  $\theta$  solving

$$J(\theta) = \frac{1}{n_0 + n_1} \left\{ \sum_{i=1}^{n_0} \left( -1 - u(X_i^0, \theta) \right)^2 + \sum_{j=1}^{n_1} \left( 1 - u(X_j^1, \theta) \right)^2 \right\}$$

$$\min_{\theta} J(\theta) \Rightarrow \theta_0 \Rightarrow u(X, \theta_0)$$

For every  $X$  to test decide as follows:  $u(X, \theta_0) \underset{H_0}{\overset{H_1}{\gtrless}} 0$

Works “well”!! Why??

## Understanding using Asymptotic Analysis

$$n_0, n_1 \rightarrow \infty, \quad u(X, \theta) \rightarrow v(X)$$

$$J(\theta) = \frac{n_0}{n_0 + n_1} \frac{1}{n_0} \sum_{i=1}^{n_0} \left(1 + u(X_i^0, \theta)\right)^2 + \frac{n_1}{n_0 + n_1} \frac{1}{n_1} \sum_{j=1}^{n_1} \left(1 - u(X_j^1, \theta)\right)^2$$

$$J(v) = \mathbb{P}(H_0) \mathbb{E}_0 \left[ \left(1 + v(X)\right)^2 \right] + \mathbb{P}(H_1) \mathbb{E}_1 \left[ \left(1 - v(X)\right)^2 \right]$$

$$\min_{\theta} J(\theta) \rightarrow \min_v J(v)$$

$$\theta_o \Rightarrow u(X, \theta_o) \approx v_o(X)$$



$$\mathbb{E}_1 \left[ \left(1 - v(X)\right)^2 \right] = \mathbb{E}_0 \left[ \left(1 - v(X)\right)^2 \frac{f_1(X)}{f_0(X)} \right]$$

$$J(v) = \mathbb{P}(H_0) \mathbb{E}_0 \left[ \left(1 + v(X)\right)^2 + r(X) \left(1 - v(X)\right)^2 \right] \quad r(X) = \frac{f_1(X) \mathbb{P}(H_1)}{f_0(X) \mathbb{P}(H_0)}$$

minimize for each  $X$

$$v_o(X) = \frac{r(X) - 1}{r(X) + 1} = \omega(r(X)), \quad \text{where } \omega(r) = \frac{r - 1}{r + 1} \text{ strictly increasing}$$

Test equivalent to Bayes:  $v_o(X) = \omega(r(X)) \underset{H_0}{\overset{H_1}{\geq}} \omega(1) = 0 \Rightarrow u(X, \theta_o) \underset{H_0}{\overset{H_1}{\geq}} 0$

Equivalence in the limit

**Consistency** (with respect to the Bayes test)

$$\min_{\theta} J(\theta) = \min_{\theta} \left\{ \sum_{i=1}^{n_0} \left(1 + u(X_i^0, \theta)\right)^2 + \sum_{j=1}^{n_1} \left(1 - u(X_j^1, \theta)\right)^2 \right\}$$

$$\Rightarrow \theta_0 \Rightarrow u(X, \theta_0)$$

$$\min_{\nu} J(\nu) = \min_{\nu} \left\{ \mathbb{P}(H_0) \mathbb{E}_0 \left[ \left(1 + \nu(X)\right)^2 \right] + \mathbb{P}(H_1) \mathbb{E}_1 \left[ \left(1 - \nu(X)\right)^2 \right] \right\}$$

$$\Rightarrow \nu_0(X) = \omega(r(X))$$

Expect:  $u(X, \theta_0) \approx \omega(r(X))$

Optimum Test:  $\omega(r(X)) \underset{H_0}{\overset{H_1}{\gtrless}} \omega(1)$ , Close to Optimum:  $u(X, \theta_0) \underset{H_0}{\overset{H_1}{\gtrless}} \omega(1)$

Develop data driven methods for estimation of  $\omega(r(X))$  for other  $\omega(r)$

Consistent tests **eventually** prevail over inconsistent tests

# Likelihood Ratio Estimation

$$r(X) \underset{H_0}{\overset{H_1}{\gtrless}} 1 \equiv \omega_1(r(X)) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_1(1) \equiv \omega_2(r(X)) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_2(1)$$

$$u(X, \theta_0) \underset{H_0}{\overset{H_1}{\gtrless}} 1 \not\equiv u_1(X, \theta_1) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_1(1) \not\equiv u_2(X, \theta_2) \underset{H_0}{\overset{H_1}{\gtrless}} \omega_2(1)$$

For function  $\omega(r)$  can we define cost

$$J(\mathbf{v}) = \mathbb{P}(H_0)\mathbb{E}_0 [\phi(\mathbf{v}(X))] + \mathbb{P}(H_1)\mathbb{E}_1 [\psi(\mathbf{v}(X))]$$

so that  $\min_{\mathbf{v}} J(\mathbf{v}) \Rightarrow \mathbf{v}_o(X) = \omega(r(X))$  ?

THEOREM: Select **strictly increasing** function  $\omega(r)$  and **strictly negative** function  $\rho(z)$ . Define

$$\psi'(z) = \rho(z), \quad \phi'(z) = -\omega^{-1}(z)\rho(z)$$

then the solution of the optimization problem

$$\min_{\mathbf{v}} J(\mathbf{v}) = \min_{\mathbf{v}} \left\{ \mathbb{P}(\mathbf{H}_0) \mathbb{E}_0 [\phi(\mathbf{v}(X))] + \mathbb{P}(\mathbf{H}_1) \mathbb{E}_1 [\psi(\mathbf{v}(X))] \right\}$$

satisfies  $\mathbf{v}_0(X) = \arg \min_{\mathbf{v}} J(\mathbf{v}) = \omega(r(X))$

**Same** optimal solution  $\omega(r(X))$  for all functions  $\rho(z) < 0$

## Examples of functions

A:  $\omega(r) = r \in \mathbb{R}_+$  (likelihood ratio)

$$\rho(z) = -1, z \geq 0 \Rightarrow \phi(z) = \frac{z^2}{2}, \psi(z) = -z$$

Mean  
Square

B:  $\omega(r) = \log(r) \in \mathbb{R}$  (log-likelihood ratio)

$$\rho(z) = -e^{-0.5z} \Rightarrow \phi(z) = 2e^{0.5z}, \psi(z) = 2e^{-0.5z}$$

Exponential

C:  $\omega(r) = \frac{r}{r+1} \in [0, 1]$  (posterior probability)

$$\rho(z) = -\frac{1}{z}, z \in [0, 1] \Rightarrow \phi(z) = -\log(1-z), \psi(z) = -\log(z)$$

Cross  
Entropy

## Data Driven Implementation

$$J(\mathbf{v}) = \mathbb{P}(\mathbf{H}_0)\mathbb{E}_0 [\phi(\mathbf{v}(X))] + \mathbb{P}(\mathbf{H}_1)\mathbb{E}_1 [\psi(\mathbf{v}(X))]$$

$$J(\theta) = \frac{1}{n_0 + n_1} \left\{ \sum_{i=1}^{n_0} \phi(\mathbf{u}(X_i^0, \theta)) + \sum_{j=1}^{n_1} \psi(\mathbf{u}(X_j^1, \theta)) \right\}$$
$$\mathbf{u}(X, \theta_0) \approx \omega \left( \frac{f_1(X)\mathbb{P}(\mathbf{H}_1)}{f_0(X)\mathbb{P}(\mathbf{H}_0)} \right)$$

$$J(\theta) = \frac{1}{n_0} \sum_{i=1}^{n_0} \phi(\mathbf{u}(X_i^0, \theta)) + \frac{1}{n_1} \sum_{j=1}^{n_1} \psi(\mathbf{u}(X_j^1, \theta))$$
$$\mathbf{u}(X, \theta_0) \approx \omega \left( \frac{f_1(X)}{f_0(X)} \right)$$

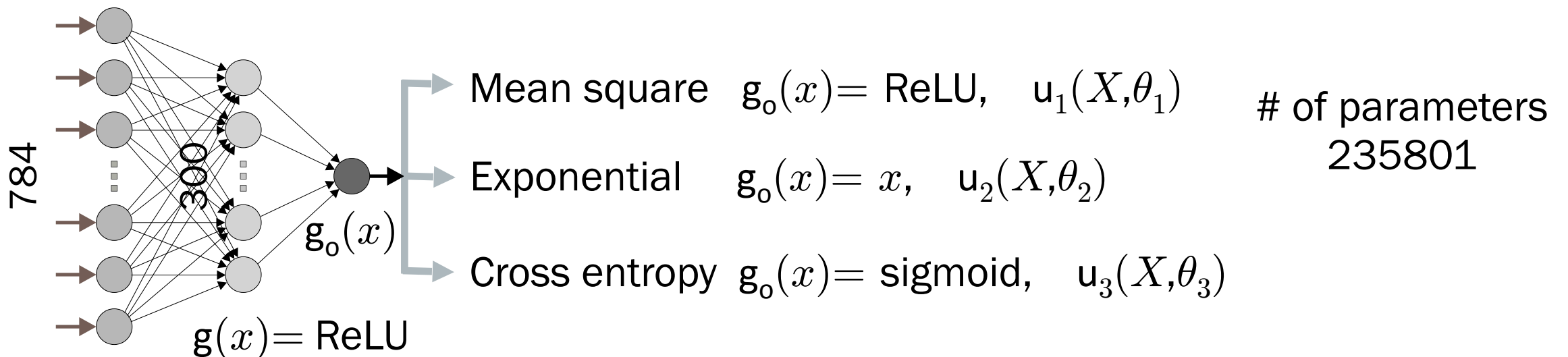
## Example: Classification Problem

From dataset MNIST isolate handwritten numerals 4 and 9



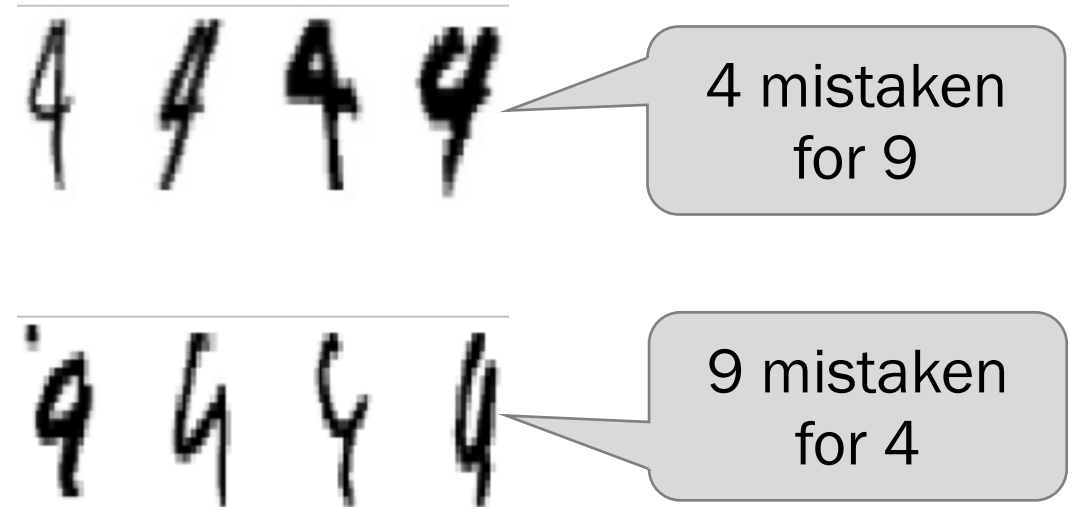
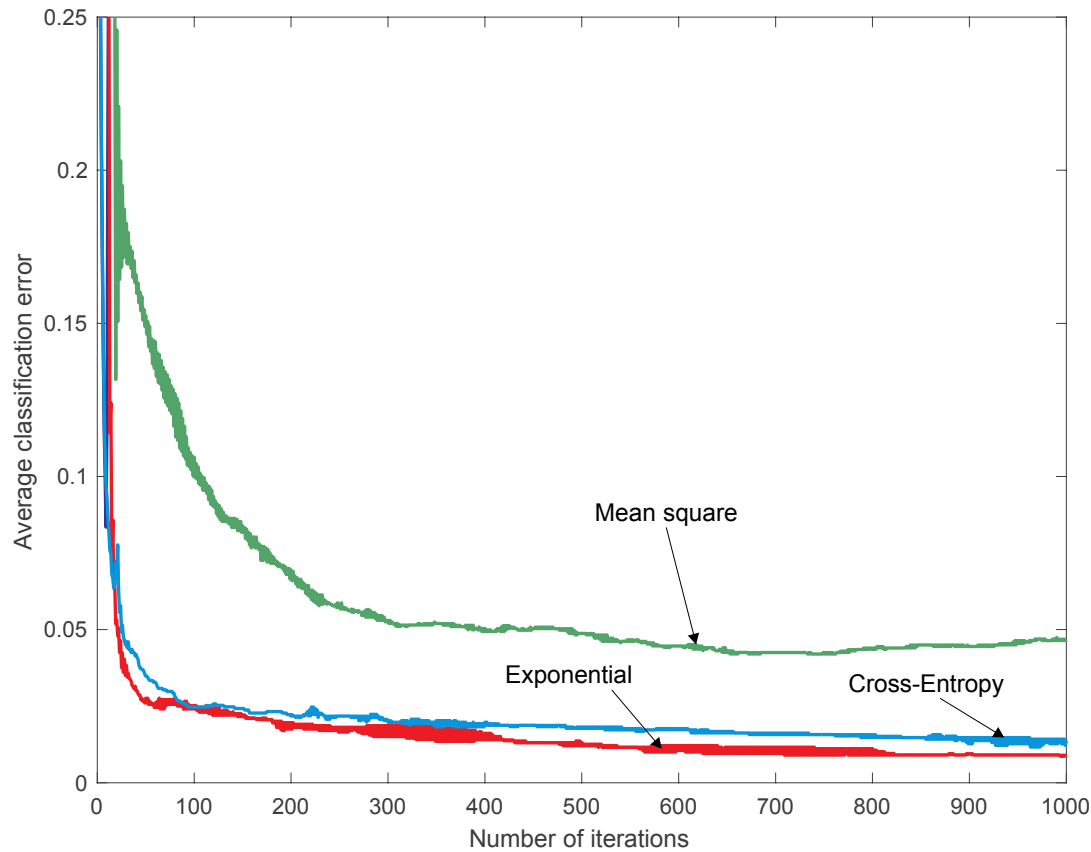
Gray scale images  $28 \times 28 = 784$  pixels. Design classifier using training data. Examine performance using testing data.

Neural network  $784 \times 300 \times 1$



Training set: 5500 “4” and 5500 “9”. Testing set: 982 “4” and 1009 “9”

$$u_1(X, \theta_1) \underset{H_0}{\overset{H_1}{\gg}} 1, \quad u_2(X, \theta_2) \underset{H_0}{\overset{H_1}{\gg}} 0, \quad u_3(X, \theta_3) \underset{H_0}{\overset{H_1}{\gg}} \frac{1}{2}$$





# Detection in Time Series

More practically interesting case: Testing of time series  $\{ X_1, X_2, \dots, X_n \}$

The **whole** set of measurements under  $H_0$  or  $H_1$

For testing we need likelihood ratio

$$\frac{f_1(X_n, \dots, X_1)}{f_0(X_n, \dots, X_1)} = \frac{f_1(X_n | X_{n-1}, \dots, X_1)}{f_0(X_n | X_{n-1}, \dots, X_1)} \frac{f_1(X_{n-1} | X_{n-2}, \dots, X_1)}{f_0(X_{n-1} | X_{n-2}, \dots, X_1)} \dots \frac{f_1(X_1)}{f_0(X_1)}$$

When i.i.d. under each hypothesis

$$\frac{f_1(X_n, \dots, X_1)}{f_0(X_n, \dots, X_1)} = \frac{f_1(X_n)}{f_0(X_n)} \dots \frac{f_1(X_1)}{f_0(X_1)}$$

Test to be used

$$\sum_{i=1}^n \log \left( \frac{f_1(X_i)}{f_0(X_i)} \right) \underset{H_0}{\overset{H_1}{\gtrless}} \eta$$

Interested in estimating  $\omega(r(X)) = \omega\left(\frac{f_1(X)}{f_0(X)}\right)$

We are given training data:  $\{X_1^0, \dots, X_{n_0}^0\}$  following  $H_0$   
 $\{X_1^1, \dots, X_{n_1}^1\}$  following  $H_1$

For each  $\omega(r)$  of interest, minimize corresponding  $J(\theta)$

$$J(\theta) = \frac{1}{n_0} \sum_{i=1}^{n_0} \phi(u(X_i^0, \theta)) + \frac{1}{n_1} \sum_{j=1}^{n_1} \psi(u(X_j^1, \theta))$$

$$u(X, \theta_0) \approx \omega\left(\frac{f_1(X)}{f_0(X)}\right)$$

For  $\omega(r) = r$ ,  $u_1(X, \theta_1) \approx \frac{f_1(X)}{f_0(X)}$ , use  $\log(u_1(X, \theta_1))$  (Mean Square)

For  $\omega(r) = \log(r)$ ,  $u_2(X, \theta_2) \approx \log\left(\frac{f_1(X)}{f_0(X)}\right)$ , use  $u_2(X, \theta_2)$  (Exponential)

For  $\omega(r) = \frac{r}{r+1}$ ,  $u_3(X, \theta_3) \approx \frac{\frac{f_1(X)}{f_0(X)}}{\frac{f_1(X)}{f_0(X)} + 1}$ , use  $\log\left(\frac{u_3(X, \theta_3)}{1 - u_3(X, \theta_3)}\right)$   
(Cross Entropy)

## Example: Testing i.i.d. sequences

$X_i$  length 10,  $f_0(X) \sim \mathcal{N}(0, I)$

$f_1(X) \sim \mathcal{N}(\frac{1}{\sqrt{10}}[1 \cdots 1]^T, 1.2I)$

Neural Network  $10 \times 20 \times 1$

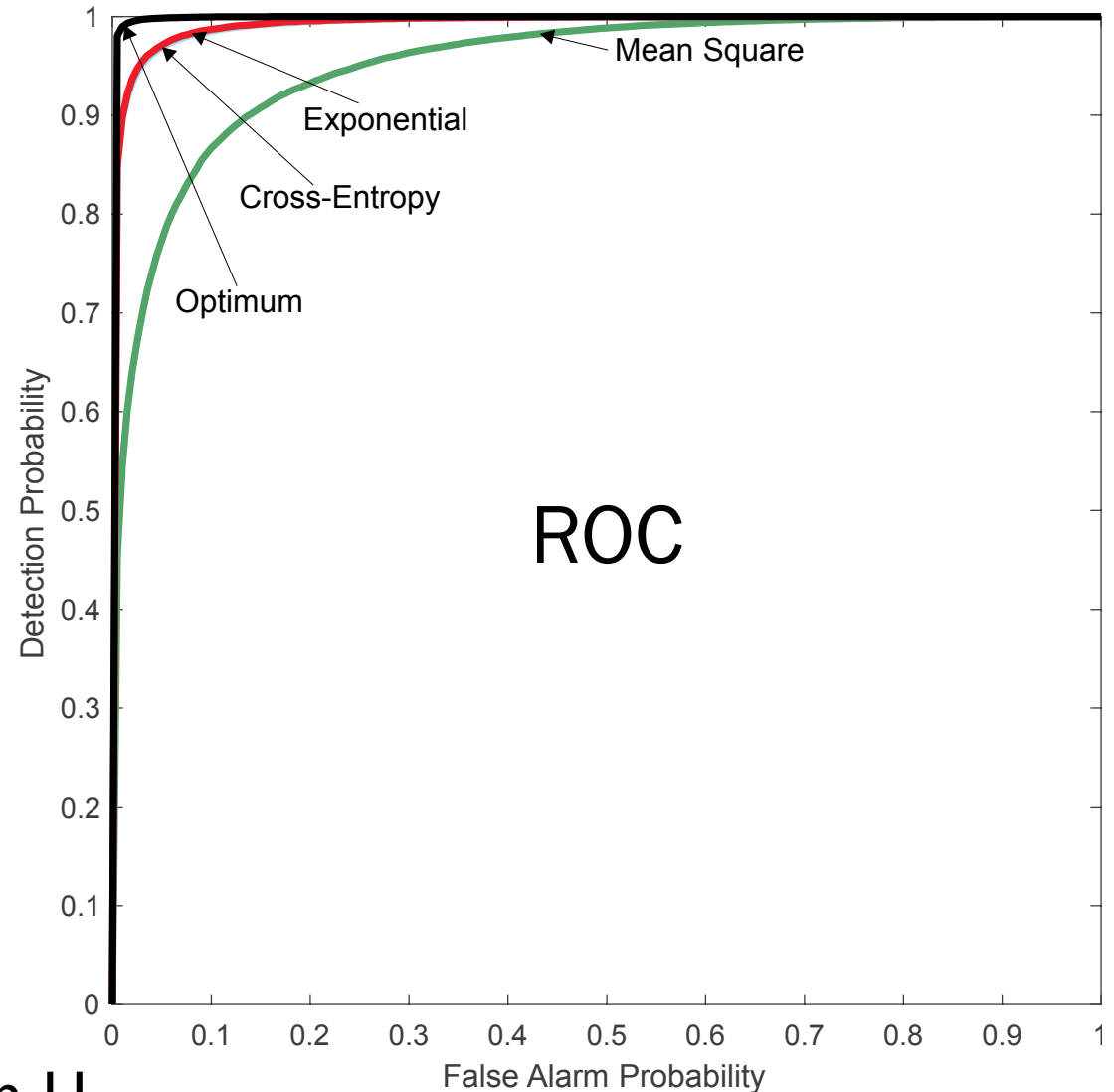
Training data  $n_0 = n_1 = 100$

Produce  $u_1(X, \theta_1), u_2(X, \theta_2), u_3(X, \theta_3)$

Testing for  $n = 20$  samples

$$\left. \begin{array}{l} \sum_{i=1}^{20} \log(u_1(X_i, \theta_1)) \\ \sum_{i=1}^{20} u_2(X_i, \theta_2) \\ \sum_{i=1}^{20} \log\left(\frac{u_3(X_i, \theta_3)}{1-u_3(X_i, \theta_3)}\right) \end{array} \right\} \begin{array}{l} H_1 \\ \text{---} \\ H_0 \end{array} \eta$$

$100000 \times 20$  realizations from  $H_0$  and from  $H_1$



# Markovian processes

Consider Markovian processes with “memory”  $m$

$$\frac{f_1(X_n | X_{n-1}, \dots, X_1)}{f_0(X_n | X_{n-1}, \dots, X_1)} = \frac{f_1(X_n | X_{n-1}, \dots, X_{n-m})}{f_0(X_n | X_{n-1}, \dots, X_{n-m})}$$

$$\frac{f_1(X_n, \dots, X_1)}{f_0(X_n, \dots, X_1)} = \frac{f_1(X_n | X_{n-1}, \dots, X_{n-m})}{f_0(X_n | X_{n-1}, \dots, X_{n-m})} \dots \frac{f_1(X_{m+1} | X_m, \dots, X_1)}{f_0(X_{m+1} | X_m, \dots, X_1)} \times \frac{f_1(X_m, \dots, X_1)}{f_0(X_m, \dots, X_1)}$$

Can we estimate likelihood ratio of conditional densities?

- a) Through data dynamics (classical)
- b) Directly (proposed)

# Classical Approach

Most common model, Autoregressive

$$X_t = A_1^i X_{t-1} + \dots + A_m^i X_{t-m} + W_t, \quad i = 0, 1$$

$$X_t = G_i(X_{t-1}, \dots, X_{t-m}, \theta^i) + W_t$$

Use training data  $\{X_1^0, \dots, X_{n_0}^0\}$  and  $\{X_1^1, \dots, X_{n_1}^1\}$  to solve

$$\min_{\theta^i} \sum_{t=1}^{n_i} \left( X_t^i - G_i(X_{t-1}^i, \dots, X_{t-m}^i, \theta^i) \right)^2 \Rightarrow \theta_0^i$$

$$W_t^i = X_t^i - G_i(X_{t-1}^i, \dots, X_{t-m}^i, \theta_0^i), \quad \Sigma_i = \frac{1}{n_i} \sum_{t=1}^{n_i} W_t^i (W_t^i)^\top$$

Assume  $\{W_t^i\}$  i.i.d. Gaussian  $\mathcal{N}(0, \Sigma_i)$

$X_t$  given  $\{X_{t-1}, \dots, X_{t-m}\}$  under hypothesis  $H_i$

Gaussian with mean  $G_i(X_{t-1}, \dots, X_{t-m}, \theta_0^i)$  and covariance  $\Sigma_i$

To test  $\{X_1, \dots, X_n\}$

$$W_t^i = X_t - G_i(X_{t-1}, \dots, X_{t-m}, \theta_0^i)$$

$$\frac{f_1(X_t | X_{t-1}, \dots, X_{t-m})}{f_0(X_t | X_{t-1}, \dots, X_{t-m})} = \frac{e^{-\frac{1}{2}(W_t^1)^\top \Sigma_1^{-1} W_t^1}}{e^{-\frac{1}{2}(W_t^0)^\top \Sigma_0^{-1} W_t^0}} \sqrt{\frac{|\Sigma_0|}{|\Sigma_1|}}$$

**Not purely data driven**

Gaussian assumption arbitrary, not necessarily suitable for all data!

## Proposed Approach

$$\log \left( \frac{f_1(X_t | X_{t-1}, \dots, X_{t-m})}{f_0(X_t | X_{t-1}, \dots, X_{t-m})} \right) = \log \left( \frac{\frac{f_1(X_t, X_{t-1}, \dots, X_{t-m})}{f_1(X_{t-1}, \dots, X_{t-m})}}{\frac{f_0(X_t, X_{t-1}, \dots, X_{t-m})}{f_0(X_{t-1}, \dots, X_{t-m})}} \right) =$$

$$\log \left( \frac{f_1(X_t, X_{t-1}, \dots, X_{t-m})}{f_0(X_t, X_{t-1}, \dots, X_{t-m})} \right)$$

$$u_{m+1}(X_t, \dots, X_{t-m}, \theta_{m+1})$$

$$- \log \left( \frac{f_1(X_{t-1}, \dots, X_{t-m})}{f_0(X_{t-1}, \dots, X_{t-m})} \right)$$

$$u_m(X_{t-1}, \dots, X_{t-m}, \theta_m)$$

$$\log \left( \frac{f_1(X_t | X_{t-1}, \dots, X_{t-m})}{f_0(X_t | X_{t-1}, \dots, X_{t-m})} \right) \approx$$

$$u_{m+1}(X_t, \dots, X_{t-m}, \theta_{m+1}) - u_m(X_{t-1}, \dots, X_{t-m}, \theta_m)$$



# Example: Testing Markov sequences (proof of concept)

Scalar observations  $\{x_1, \dots, x_n\}$

$w_t \sim \mathcal{N}(0, 1)$ , i.i.d.

$H_0 : x_t = w_t$

$H_1 : x_t = \text{sign}(x_{t-1})\sqrt{|x_{t-1}|} + w_t$

$u_2(x_t, x_{t-1}, \theta_2) : 2 \times 20 \times 1$

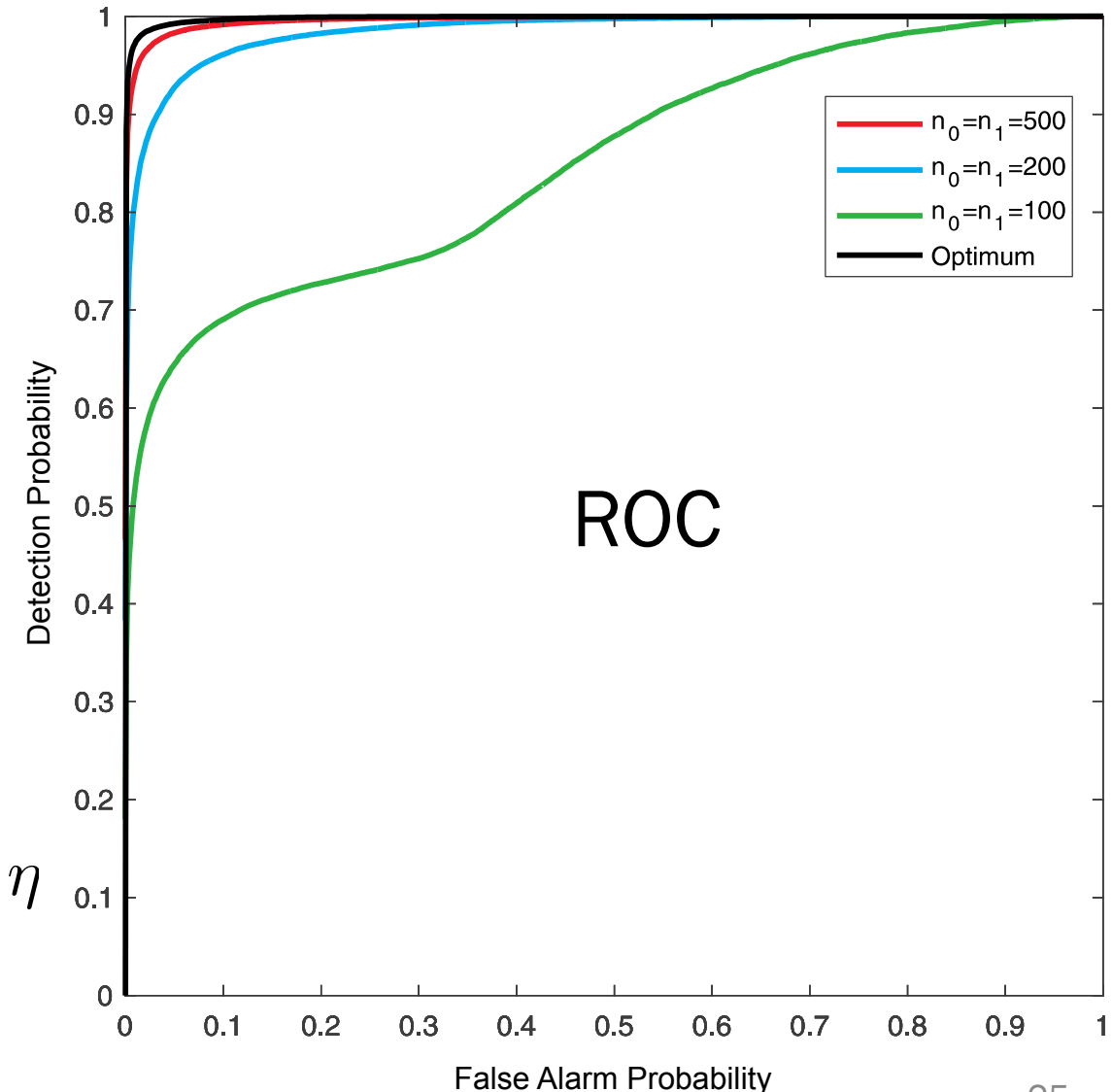
$u_1(x_t, \theta_1) : 1 \times 10 \times 1$  (Exponential)

Training data  $n_0 = n_1 = 100, 200, 500$

Testing  $n = 20$ , i.e.  $\{x_1, \dots, x_{20}\}$

$\sum_{t=2}^{20} u_2(x_t, x_{t-1}, \theta_2) - \sum_{t=2}^{19} u_1(x_t, \theta_1) \underset{H_0}{\overset{H_1}{\geq}} \eta$

100000  $\times$  20 samples from  $H_0$  and  $H_1$



# What is Missing from the Analysis ?

- For fixed  $\omega(r)$  find  $\rho(z)$  that produces best estimates for  $\omega(r(X))$   
Requires analysis of estimation performance of neural networks for finite models (extremely challenging !!!).
- Rank functions  $\omega(r)$  according to their approximation accuracy  
Is it possible to show that Mean Square is worse than Exponential or Cross Entropy?

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